

Lecture 6: Concentration Inequalities

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Overview

- Introduction
 - Final Project, Collaboration & Academic Integrity
 - Concentration Inequalities
 - Markov's Inequality
- Higher Moments
 - Moments and Variance
 - Chebyshev's Inequality
 - Chernoff-Hoeffding's Inequality
- Acknowledgements

Final Project

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- Probably many of you may have similar questions about the final project, so if you want to ask us something, piazza would be great so that everyone can participate in the discussion! :)
- There is a post pinned on piazza for you all to look for partners for your final project (undergraduates). So if you have a project in mind and want to check if someone else is interested in working with you on it, please post it there!

Collaboration on Homework

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- Writing proofs that are correct (or that correctly showcase your ideas) is part of you mathematical development! (as well as checking that your proof is correct)
- Solutions to the homework problems *should be simple*. So, if things are getting very complicated in your solution, there is probably another way (this is a **general hint**)

Why do we want concentration?

When evaluating performance of randomized algorithms, not enough to know our algorithm runs in expected time T . What we want to say is

“our algorithm will run in time $\approx T$ *very often*.”

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Running time *small* with *high probability better than* small expected running time.

Often times in algorithm analysis, running time is *concentrated* around expectation. This *concentration of measure* proves that our algorithms will *typically* run in time close to expectation.

Today's inequalities

Theorem (Markov's Inequality)

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$$\Pr[X \geq t] \leq \frac{\mathbb{E}[X]}{t}, \quad \forall t > 0.$$

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Theorem (Chebyshev's Inequality)

Let X be a ~~non-negative~~ discrete random variable. Then

$$\Pr[|X - \mathbb{E}[X]| \geq t] \leq \frac{\text{Var}[X]}{t^2}, \quad \forall t > 0.$$

Today's inequalities II

Theorem (Chernoff-Hoeffding's Inequality)

Let X_1, \dots, X_n be independent indicator variables such that $\Pr[X_i = 1] = p_i$, where $0 < p_i < 1$. Let $X = \sum_{i=1}^n X_i$ and $\delta > 0$. Then

$$\Pr[X \geq (1 + \delta) \cdot \mathbb{E}[X]] \leq \left[\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right]^{\mathbb{E}[X]},$$

and

$$\Pr[X \leq (1 - \delta) \cdot \mathbb{E}[X]] \leq \exp(-\mathbb{E}[X] \cdot \delta^2/2).$$

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$$\Pr[X \geq t] \leq \frac{\mathbb{E}[X]}{t}, \quad \forall t > 0.$$

Proof: $\mathbb{E}[X] = \sum_{n \geq 0} \Pr[X=n] \cdot n =$

$$= \sum_{n=0}^{t-1} \underbrace{\Pr[X=n]}_{\geq 0} \cdot n + \sum_{n \geq t} \underbrace{n}_{\geq t} \cdot \Pr[X=n]$$

$$\geq 0 + t \cdot \sum_{n \geq t} \Pr[X=n] = t \cdot \Pr[X \geq t]. \quad \square$$

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Useful when we have no information beyond expected value (or when random variable difficult to analyze). Otherwise other inequalities much sharper!

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Some practice problems.

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- Does it hold for general random variables (not just non-negative)?
- Can it be modified to upper bound $\Pr[X \leq t]$?

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Moments and Variance

To give better bounds, we need more information about the random variable (beyond expected value).

How to distinguish between:

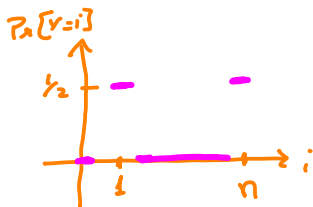
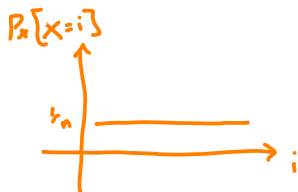
Moments and Variance

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How to distinguish between:

- X such that $\Pr[X = i] = \begin{cases} 1/n, & \text{if } 1 \leq i \leq n \\ 0, & \text{otherwise} \end{cases}$

- Y such that $\Pr[Y = 1] = 1/2$ and $\Pr[Y = n] = 1/2$ } *always far from mean*



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- same expectation, but very different random variables...

$$\underbrace{E[X]} = E[Y] = \frac{n+1}{2}$$

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- same expectation, but very different random variables...
- Look at how far variable usually is from its expectation. How to do that?

How far X is from $\mathbb{E}[X]$: $|X - \mathbb{E}[X]|$

Can try to compute
 $\mathbb{E}[|X - \mathbb{E}[X]|]$

if close to expectation \leftarrow small
if always far from expectation \leftarrow large

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Theorem (Chebyshev's Inequality)

Let X be a ~~non-negative~~ discrete random variable. Then

$$\Pr[|X - \mathbb{E}[X]| \geq t] \leq \frac{\text{Var}[X]}{t^2}, \quad \forall t > 0.$$

Chebyshev's inequality

Let X be a random variable.

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- Its **Variance** is defined as $\text{Var}[X] := \mathbb{E}[(X - \mathbb{E}[X])^2]$
- and its **standard deviation** is $\sigma(X) := \sqrt{\text{Var}[X]}$

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Theorem (Chebyshev's Inequality)

Let X be a ~~continuous~~ discrete random variable. Then

$$\Pr[|X - \mathbb{E}[X]| \geq t] \leq \frac{\text{Var}[X]}{t^2}, \quad \forall t > 0.$$

Proof: only thing we know is Markov. Let's use it!

$$Y := (X - \mathbb{E}[X])^2 \quad Y: \begin{cases} \text{discrete if } X \text{ discrete} \\ \geq 0 \end{cases} \quad (\text{can use Markov!})$$

By Markov: $\Pr[X - \mathbb{E}[X] \geq t^2] \leq \frac{\mathbb{E}[Y]}{t^2} = \frac{\text{Var}[X]}{t^2}$

and $\Pr[|X - \mathbb{E}[X]| \geq t] = \Pr[X - \mathbb{E}[X] \geq t^2] \quad \square$

Covariance

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Definition (Covariance)

The *covariance* of two random variables X, Y is defined as

$$\text{Cov}[X, Y] := \mathbb{E}[(X - \mathbb{E}[X]) \cdot (Y - \mathbb{E}[Y])].$$

We say that X, Y are *positively correlated* if $\text{Cov}[X, Y] > 0$ and *negatively correlated* if $\text{Cov}[X, Y] < 0$.

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Proposition

- $\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2 \text{Cov}[X, Y]$
- If X, Y are independent, then $\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y]$

Practice problem: prove this proposition!

Chebyshev & Covariance example

Coin Flipping: If X be # heads in n independent unbiased coin flips, let us bound again $\Pr[X \geq 3n/4]$.

$$X_i = \begin{cases} 1 & \text{if } i\text{th coin flipped heads} \\ 0 & \text{otherwise} \end{cases}$$

$$X = \sum_{i=1}^n X_i, \quad X_i, X_j \text{ independent.}$$

By proposition $\text{Var}[X] = \sum_{i=1}^n \text{Var}[X_i] = \sum_{i=1}^n \frac{1}{4} = \frac{n}{4}$

Chebyshev: $\Pr[X \geq 3n/4] \leq \Pr[|X - \underbrace{n/2}_{\mathbb{E}[X]}| \geq n/4]$

$$\leq \frac{n/4}{(n/4)^2} = \frac{4}{n}$$

□

Much better than Markov!

Higher Moments

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$$\mu_X^{(k)} := \mathbb{E}[(X - \mathbb{E}[X])^k],$$

if it exists.

$$\text{Var}[X] = \mu_X^{(2)}$$

Practice problem: when will the k^{th} moment not exist? (Appendix C of MR'07).
(will post on this on my webpage)

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Chebyshev's inequality is most useful when we only have information about the *second moment* of our random variable X .

Practice problem: Can you generalize Chebyshev's inequality to k^{th} order moments?

Sums of Independent Random Variables

Often times in analysis of algorithms we deal with random variables which are sums of independent random variables (see Distinct Elements analysis from last lecture, hashing, etc). *balls & bins lecture!*

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Law of large numbers: average of independent, identically distributed variables is approximately the expectation of the random variables. That is, if each X_i is an independent copy of random variable X

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Central Limit Theorem: if we let $Z_n = \sum_{i=1}^n X_i$, where X_i independent copy of X , the random variable

$$Y_n = \frac{Z_n - n \cdot \mathbb{E}[X]}{\sqrt{n \cdot \sigma(X)^2}} \rightarrow \mathcal{N}(0, 1)$$

Gaussian distribution

Chernoff Bounds

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Simple Setting: we have n coin flips, each is head with probability p . So

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- Not easy to work with, hard to generalize (homework 1 question 6)

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Chernoff Bounds

Generic Chernoff Bounds: apply Markov in the following way:

$$\Pr[X \geq a] = \Pr[e^{tX} \geq e^{ta}] \leq \mathbb{E}[e^{tX}] / e^{ta}, \quad \text{for any } t > 0.$$

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What do we gain by doing this?

- If $X = X_1 + X_2$, where X_1, X_2 are independent, note that

$$\mathbb{E}[e^{tX}] = \mathbb{E}[e^{tX_1} e^{tX_2}] = \mathbb{E}[e^{tX_1}] \cdot \mathbb{E}[e^{tX_2}]$$

easy to bound expectation by splitting it
(yes, but we know of linearity of expectation...)

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$$\mathbb{E}[e^{tX}] = \mathbb{E}[e^{tX_1} e^{tX_2}] = \mathbb{E}[e^{tX_1}] \cdot \mathbb{E}[e^{tX_2}]$$

- The *moment generating function*

$$M_X(t) := \mathbb{E}[e^{tX}] = \mathbb{E} \left[\sum_{i \geq 0} \frac{t^i}{i!} \cdot X^i \right] = \sum_{i \geq 0} \frac{t^i}{i!} \cdot \mathbb{E}[X^i]$$

contains information about all moments!

so expect to be more powerful than Chebyshev (2^{nd} moment)

Chernoff Bounds for Bounded Variables

Example (Heterogeneous Coin Flips)

Let $X_i = \begin{cases} 1, & \text{with probability } p_i \\ 0, & \text{otherwise} \end{cases}$, $X = \sum_{i=1}^n X_i$ and $\mu = \mathbb{E}[X]$

① for $\delta > 0$, $\Pr[X \geq (1 + \delta)\mu] \leq \left[\frac{e^\delta}{(1 + \delta)^{1 + \delta}} \right]^\mu$

$$\mu = \mathbb{E}[X] = \sum_{i=1}^n \mathbb{E}[X_i] = \sum_{i=1}^n p_i$$

Proof: $\Pr[X \geq (1 + \delta)\mu] = \Pr[X e^{-t(1 + \delta)\mu} \geq e^{-t(1 + \delta)\mu} e^{t(1 + \delta)\mu}] \leq \mathbb{E}[e^{tX}] / e^{t(1 + \delta)\mu}$

$$= \frac{1}{e^{t(1 + \delta)\mu}} \cdot \prod_{i=1}^n \mathbb{E}[e^{tX_i}] = \frac{1}{e^{t(1 + \delta)\mu}} \cdot \prod_{i=1}^n (e^t p_i + (1 - p_i)) \leq$$

$$\leq \frac{1}{e^{t(1 + \delta)\mu}} \cdot \prod_{i=1}^n e^{p_i(e^t - 1)} = \frac{1}{e^{t(1 + \delta)\mu}} \cdot e^{(e^t - 1)\sum p_i} = \left(\frac{e^{e^t - 1}}{e^{t(1 + \delta)}} \right)^\mu \quad t = \ln(1 + \delta)$$

Chernoff Bounds for Bounded Variables

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① for $\delta > 0$, $\Pr[X \geq (1 + \delta)\mu] \leq \left[\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right]^\mu$

② for $0 < \delta < 1$, $\Pr[X \geq (1 + \delta)\mu] \leq e^{-\delta^2\mu/3}$

just note that $0 < \delta < 1 \Rightarrow \frac{e^\delta}{(1+\delta)^{1+\delta}} < e^{-\delta^2/3}$

Then consider $f(\delta) = \delta - (1+\delta)\ln(1+\delta) + \delta^2/3$ and

show $f(\delta) \leq 0$ in $[0,1]$.

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Example (Heterogeneous Coin Flips)

Let $X_i = \begin{cases} 1, & \text{with probability } p_i \\ 0, & \text{otherwise} \end{cases}$, $X = \sum_{i=1}^n X_i$ and $\mu = \mathbb{E}[X]$

- 1 for $\delta > 0$, $\Pr[X \geq (1 + \delta)\mu] \leq \left[\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right]^\mu$
- 2 for $0 < \delta < 1$, $\Pr[X \geq (1 + \delta)\mu] \leq e^{-\delta^2\mu/3}$
- 3 for $R \geq 6\mu$, $\Pr[X \geq R] \leq 2^{-R}$

$R \geq 6\mu$ then $\delta \geq 5$ in ①

Chernoff Bounds for Bounded Variables

What about the lower tail?

²See [Motwani & Raghavan 2007, Theorem 4.2] or [Mitzenmacher & Upfal, Theorem 4.5]

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Theorem (Heterogeneous Coin Flips - lower tail)

- 1 $\Pr[X \leq (1 - \delta) \cdot \mu] \leq \left[\frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right]^\mu$
- 2 if $0 < \delta < 1$ then $\Pr[X \leq (1 - \delta) \cdot \mu] \leq e^{-\mu\delta^2/2}$

Practice problem: prove this theorem!

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Hoeffding's generalization

What if the variables X_i took values in $[a_i, b_i]$?

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Theorem (Hoeffding's Inequality)

Let X_i be independent random variables, taking values in $[a_i, b_i]$,
 $X = \sum_{i=1}^n X_i$. Then

$$\Pr[|X - \mathbb{E}[X]| \geq \ell] \leq 2 \cdot \exp\left(-\frac{2\ell^2}{\sum_{i=1}^n (b_i - a_i)^2}\right)$$

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Proof uses *Hoeffding's lemma*: $\mathbb{E}[e^{t(X_i - \mathbb{E}[X_i])}] \leq \exp\left(\frac{t^2(b_i - a_i)^2}{8}\right)$

Practice problem: prove this theorem.

Remarks

- In coin flips example from beginning of lecture, by flipping n independent fair coins, expected # heads is $n/2$. Chernoff-Hoeffding implies:

$$\Pr[|\# \text{ heads} - \mu| \geq \delta\mu] \leq 2 \exp(-\mu\delta^2/3) = 2 \exp(-n\delta^2/6)$$

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- With high probability, # heads is within $O(\sqrt{n})$ of the expected value (this comes up in many places). **Practice problem:** prove that with constant probability that $|\# \text{ heads} - n/2| = \Omega(n)$.
- Recall from previous slides that Markov gave us that $\Pr[\# \text{ heads} \geq 3n/4] \leq 2/3$, and Chebyshev gave us $\Pr[\# \text{ heads} \geq 3n/4] \leq 4/n$. Chernoff gives us $\Pr[\# \text{ heads} \geq 3n/4] \leq e^{-n/24}$.

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- Why do we want to compute moments? See Sum-of-Squares and pseudo-distributions references in course webpage. These methods give very powerful tools to solve many challenging problems! (great final project topic!)
- Chernoff-Hoeffding bounds also hold for *negatively correlated* variables, because all we need is

$$\mathbb{E}[e^{t(X+Y)}] \leq \mathbb{E}[e^{tX}] \cdot \mathbb{E}[e^{tY}]$$

This observation is very useful in many applications (also great source of final projects!)

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- For instance: two edges appear in a random spanning tree is a negatively correlated event, thus Chernoff bounds are useful to analyze random spanning trees.

Acknowledgement

- Lecture based largely on Lap Chi's notes and [Motwani & Raghavan 2007, Chapters 3 and 4].
- See Lap Chi's notes at <https://cs.uwaterloo.ca/~lapchi/cs466/notes/L02.pdf>

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