

# Lecture 11: Markov Chains, Random Walks, Mixing Time, Page Rank

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# Overview

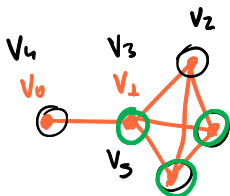
- Introduction
  - Why Random Walks & Markov Chains?
  - Basics on Theory of (finite) Markov Chains
- Main Topics
  - Fundamental Theorem of Markov Chains
  - Page Rank
- Conclusion
- Acknowledgements

# What is a Random Walk?

Given a graph  $G(V, E)$

- 1 random walk starts from a vertex  $v_0$
- 2 at each time step it moves to a *uniformly random neighbor* of the current vertex in the graph

$$v_{t+1} \leftarrow \underbrace{R}_{\text{neighbors of current vertex}} N_G(v_t)$$



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Basic questions involving random walks:

# What is a Random Walk?

Given a graph  $G(V, E)$

$$|V| = n \quad |E| = m$$

- 1 random walk starts from a vertex  $v_0 (= 1)$
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$$v_{t+1} \leftarrow_R N_G(v_t)$$

Basic questions involving random walks:

- *Stationary distribution*: does the random walk converge to a “stable” distribution? If it does, what is this distribution?

$P_t \in \mathbb{R}_{\geq 0}^n$  probability distribution over  $V$

$$N_G(1) = \{2, 3\}$$

$$P_0 = (1, 0, 0, \dots, 0)$$

$$P_1 = (0, 1/2, 1/2, 0, \dots, 0)$$

$$\{P_t\}_{t \geq 0} \rightarrow \pi?$$

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→ today's lecture

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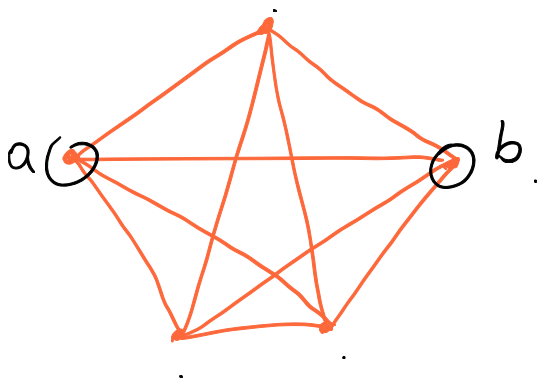
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- *Mixing time*: how long does it take for the walk to converge to the stationary distribution?
- *Hitting time*: starting from a vertex  $v_0$ , what is expected number of steps until it reaches a vertex  $v_f$ ?
- *Cover time*: how long does it take to reach every vertex of the graph at least once?



## Random Walk: Example

- Suppose  $G(V, E) = K_n$ , the complete graph,  $a, b \in V$  two vertices



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- Suppose  $G(V, E) = K_n$ , the complete graph,  $a, b \in V$  two vertices
  - 1 What is expected number of steps to reach  $b$  in simple random walk starting at  $a$ ? (i.e., hitting time)

$\pi_t = P_x[\text{reach } b \text{ first time at time } t \text{ starting at } a]$

$$\pi_t = \underbrace{\left(\frac{n-2}{n-1}\right)^{t-1}}_{\text{not reaching } b \text{ in first } t-1 \text{ steps}} \cdot \frac{1}{n-1} \underbrace{\text{reaching } b \text{ at } t^{\text{th}} \text{ step}}$$

$$E[\# \text{ steps}] = \sum_{t \geq 1} \underbrace{t}_{\# \text{ steps}} \cdot \pi_t = \sum_{t \geq 1} \frac{t}{n-1} \cdot \left(\frac{n-2}{n-1}\right)^{t-1} = \frac{1}{n-1} \sum_{t \geq 1} t \cdot \left(\frac{n-2}{n-1}\right)^{t-1}$$

$$= \frac{1}{n-1} \frac{d}{dx} \left( \sum_{t \geq 1} x^t \right) \Big|_{x = \frac{n-2}{n-1}} = \frac{1}{n-1} \frac{d}{dx} \left( \frac{x}{1-x} \right) \Big|_{x = \frac{n-2}{n-1}} = \frac{1}{n-1} \left( \frac{1}{(1-x)^2} \right) \Big|_{x = \frac{n-2}{n-1}}$$

$$= \frac{1}{n-1} \cdot (n-1)^2 = \boxed{n-1}$$

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Practice problem

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  - 3 Stationary Distribution?

$$\pi = (\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n})$$

$$\Pr[a \text{ in next step}] = \sum_{i=1}^n \Pr[\text{currently at } i] \cdot \Pr[i \rightarrow a]$$

$$= \sum_{i=1}^n \frac{1}{n} \cdot \left( \frac{1}{n-1} \cdot (1 - \delta_{ia}) \right) = \frac{1}{n} \cdot \frac{n-1}{n-1} = \frac{1}{n}$$

$$\frac{1}{n-1} \text{ if } i \neq a \\ 0 \text{ if } i = a$$

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- **Practice question:** Compare question 2 to coupon collector problem!


$$(n-1) H_{n-1}$$

harmonic number

# What is a Markov Chain?

history of random walk

Random walk is a special kind of *stochastic process*:

$$\Pr[X_t = v_t \mid X_0 = v_0, \dots, X_{t-1} = v_{t-1}] = \Pr[X_t = v_t \mid X_{t-1} = v_{t-1}]$$

in step  $t$   
I am at  
vertex  $x_t$

only useful  
information is  
the vertex at  
time  $t-1$

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Process is “*forgetful*”

*Markov chain* is characterized by this property.

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  - Permanent of non-negative matrices [Jerrum, Vigoda & Sinclair] (*great final project topic!*)

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# Representing Finite Markov Chains

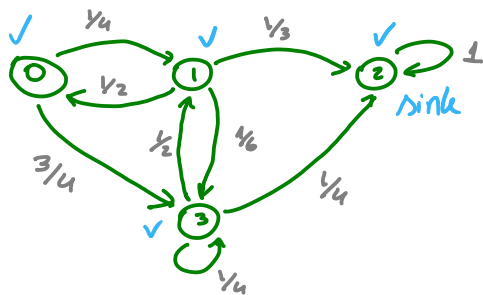
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- Vertex is a state of Markov chain
- edge  $(i, j)$  corresponds to transition probability from  $i$  to  $j$

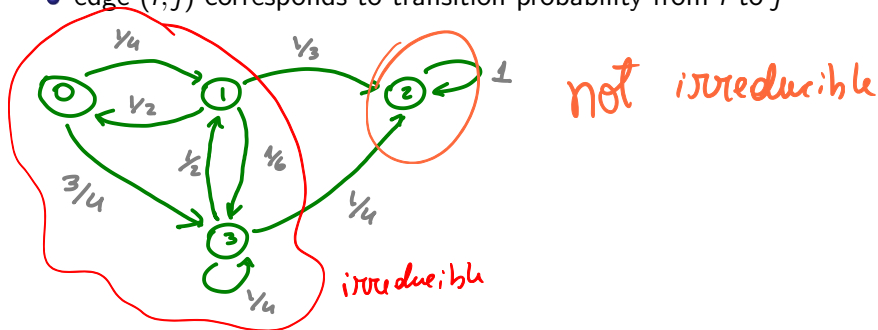


sum weights  
coming out of  
node = 1

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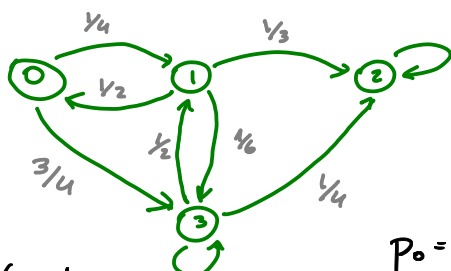
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- Markov Chain *irreducible* if underlying directed graph is *strongly connected* (i.e. there is directed path from  $i$  to  $j$  for any pair  $i, j \in V$ )

# Representing Finite Markov Chains

Markov chain can be seen in weighted adjacency matrix format.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1/4 & 0 & 3/4 \\ 0 & 0 & 1 & 0 \\ 0 & 1/2 & 1/4 & 1/4 \end{pmatrix} \end{matrix}$$

$$(0, 1/4, 0, 3/4)$$

$$p_0 = (1, 0, 0, 0)$$

$$p_1 = p_0 \cdot P = (0, 1/4, 0, 3/4)$$

- $P \in \mathbb{R}^{n \times n}$  transition matrix

each row sums to 1 (and entries are  $\geq 0$ )

- entry  $P_{ij}$  corresponds to transition probability from  $i$  to  $j$

- $p_t \in \mathbb{R}^n$  probability vector:  $p_t(i) := \Pr[\text{being at state } i \text{ at time } t]$

- Transition given by

row vector

$$p_{t+1} = p_t \cdot P$$

# Properties of Markov Chains

- *Period* of a state  $i$  is:

$$\gcd\{t \in \mathbb{N} \mid P_{i,i}^t > 0\}$$

That is, gcd of all times  $t$  such that the probability of starting at state  $i$  and being back at  $i$  at time  $t$  is positive

$P_{ii}^1$  probability that I stay at  $i$

$P_{ii}^2$  probability that I  $i \rightarrow v \rightarrow i$   
am at  $i$  after 2 steps



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## Lemma

For any *finite, irreducible and aperiodic* Markov Chain, there exists  $T < \infty$  such that

$$P_{i,j}^t > 0 \text{ for any } i, j \in V \text{ and } t \geq T.$$

That is: at some point we will reach every vertex + positive probability of being in each vertex after walking long enough!



# Stationary Distributions

## Definition (Stationary Distribution)

A stationary distribution of a Markov Chain is a probability distribution  $\pi \in \mathbb{R}^n$  such that

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- Given two distributions  $p, q \in \mathbb{R}^n$ , their *total variational distance* is

$$\Delta_{TV}(p, q) = \frac{1}{2} \sum_{i=1}^n |p_i - q_i| = \frac{1}{2} \cdot \|p - q\|_1$$

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- $p_t$  *converges* to  $q$  iff  $\lim_{t \rightarrow \infty} \Delta_{TV}(p_t, q) = 0$

# Mixing Time of Markov Chains

## Definition (Mixing Time)

The  $\varepsilon$ -mixing time of a Markov Chain is the smallest  $t$  such that

$$\Delta_{TV}(p_t, \pi) \leq \varepsilon$$

regardless of the initial starting distribution  $p_0$ .

For complete graph  $K_n$ :

transition matrix  $P = \begin{pmatrix} 0 & \dots & 1 \\ \vdots & \ddots & \vdots \\ 1 & \dots & 0 \end{pmatrix} \cdot \frac{1}{n-1}$

all ones matrix  
↑

eigenvalues:  $\det(tI - P) = \left(\frac{1}{n-1}\right)^n \cdot \det(\underbrace{[t + (n-1)t]I - J_n}_{\text{all ones matrix}})$

eigenvalues of  $J_n$ :  $\alpha_1 = n$ ,  $\alpha_2 = \alpha_3 = \dots = \alpha_n = 0$   
eigenvectors of  $J_n$ :  $v_1$ ,  $v_2, \dots, v_n$  (orthonormal)

$$\frac{1}{\sqrt{n}}(1, 1, \dots, 1)$$

$$\lambda_1 = 1$$

$$\lambda_2 = \dots = \lambda_n = -1/(n-1)$$

$\therefore$  eigenvalues of  $P$ :  $t + (n-1)\lambda_i = \alpha_i$   
eigenvectors of  $P$ :  $v_1, v_2, \dots, v_n$

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- Eigenvalues  $\lambda_1 = 1, \lambda_2 = \dots = \lambda_n = -1/(n-1)$ , corresponding eigenvectors  $v_1, \dots, v_n$  (orthonormal)  $v_i = (1, 1, \dots, 1) \cdot 1/\sqrt{n}$

$$\begin{aligned} p_0 \cdot P^t &= p_0 \left( \sum_{i=1}^n \lambda_i v_i^T v_i \right)^t = p_0 \cdot \sum_{i=1}^n \lambda_i^t \cdot v_i^T v_i = \\ &= \sum_{i=1}^n \lambda_i^t (p_0 v_i^T) \cdot v_i = \underbrace{(p_0 v_1^T) \cdot v_1}_{= 1/\sqrt{n} \text{ (prob. distribution)}} + \underbrace{(-1/(n-1))^t}_{\text{orthogonality}} \cdot \sum_{i=2}^n (p_0 v_i^T) \cdot v_i \end{aligned}$$

*Small  $\rightarrow 0$*   
*Small*

$$t = O\left(\log_{n-1}\left(\frac{n-1}{\varepsilon}\right)\right) \Rightarrow \Delta_{TV} \leq \varepsilon$$



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# Fundamental Theorem of Markov Chains

- The *return time* from state  $i$  to itself is defined as

$$H_{i,i} := \min\{t \geq 1 \mid X_t = i, X_0 = i\}$$

back at  $i$   
first time  
at time  $t$

started at  $i$

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## Theorem (Fundamental Theorem of Markov Chains)

Any *finite, irreducible and aperiodic* Markov Chain has the following properties:

- There exists a *unique* stationary distribution  $\pi$ , where  $\pi_i > 0$  for all  $i \in [n]$
- The sequence of distributions  $\{p_t\}_{t \geq 0}$  will converge to  $\pi$ , no matter what the initial distribution is.

3

$$\pi_i = \lim_{t \rightarrow \infty} P_{i,i}^t = \frac{1}{h_{i,i}}$$

# Fundamental Theorem of Markov Chains

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Any *finite, irreducible* and *aperiodic* Markov Chain has the following properties:

- 1 There is *unique* stationary distribution  $\pi$ , where  $\pi_i > 0$  for all  $i \in [n]$
- 2 For every distribution  $p_0 \in \mathbb{R}_{\geq 0}^n$ ,  $\lim_{t \rightarrow \infty} p_0 \cdot P^t = \pi$
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Intuition for proof of this theorem:

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Intuition for proof of this theorem:

- two random walks are “indistinguishable” after they “meet” at the *same vertex  $v$*  at a particular *time  $t$*
- By finiteness, irreducibility and aperiodicity, two walks will meet with positive probability (and thus by Markov property) become *same distribution*

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If our underlying graph is undirected:



# Fundamental Theorem of Markov Chains

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Any *finite, irreducible* and *aperiodic* Markov Chain has the following properties:

- 1 There is *unique* stationary distribution  $\pi$ , where  $\pi_i > 0$  for all  $i \in [n]$
- 2 For every distribution  $p_0 \in \mathbb{R}_{\geq 0}^n$ ,  $\lim_{t \rightarrow \infty} p_0 \cdot P^t = \pi$

$$\pi_i = \lim_{t \rightarrow \infty} P_{i,i}^t = \frac{1}{h_{i,i}}$$

If our underlying graph is undirected:

- If  $A_G$  adjacency matrix of  $G(V, E)$  and  $D = \text{diag}(d_1, d_2, \dots, d_n)$ , transition matrix:

$$D = \begin{pmatrix} d_1 & & 0 \\ & d_2 & \\ 0 & & \ddots \\ & & & d_n \end{pmatrix}$$

$$P = D^{-1} \cdot A_G$$

multiplies row  $i$  by  $\frac{1}{d_i}$   
(normalize neighbor weights)

$$P_{i+1}(j) = \sum_{i=1}^n P_{i+1}(i) P_{ij} = \sum_{i=1}^n P_{i+1}(i) \cdot A_{ij} \cdot \frac{1}{d_i} \rightarrow \text{uniform distribution over neighbors of } i$$

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- Note that in this case, easy to guess stationary distribution:

$$\pi_i = \frac{d_i}{2m}, \quad m = |E|$$

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new distribution

$$\pi_i = \frac{d_i}{2m}, \quad m = |E|$$

$(\pi P)_j = \sum_{i=1}^n \pi_i \cdot P_{ij} = \sum_{i=1}^n \frac{d_i}{2m} \cdot A_{ij} \cdot \frac{1}{d_i} = \sum_{i=1}^n \frac{A_{ij}}{2m} = \frac{d_j}{2m} = \tilde{\pi}_j$

$\frac{1}{d_i}$  is the  $j$ th entry of old distribution

$A_{ij} = 1$  iff  $i \in N_c(j)$   
0 otherwise

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- $P$  not symmetric, but *similar* to a symmetric matrix:

$$\cancel{D^{-1/2} \cdot P \cdot D^{1/2}} = \cancel{D^{-1/2} \cdot D^{-1} \cdot A_G \cdot D^{1/2}} = \cancel{D^{-1/2} \cdot A_G \cdot D^{1/2}}$$
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conjugate of  $P$  symmetric symmetric

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spectral thm

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    - ✓ • eigenvalue is *positive*
  - This eigenvector is  $\pi$ !

$$\pi P = \pi$$

$\pi$  eigenvector

$$\pi_i > 0$$

$P$  has max eigenvalue 1  
all other eigenvalues  $< 1$  in absolute value

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  - All random walks converge to  $\pi$ , as we wanted to show.

*same analysis that we did for kn*

- Introduction
  - Why Random Walks & Markov Chains?
  - Basics on Theory of (finite) Markov Chains
- Main Topics
  - Fundamental Theorem of Markov Chains
  - Page Rank
- Conclusion
- Acknowledgements

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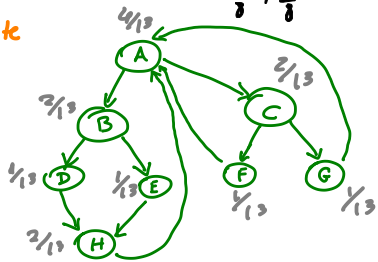
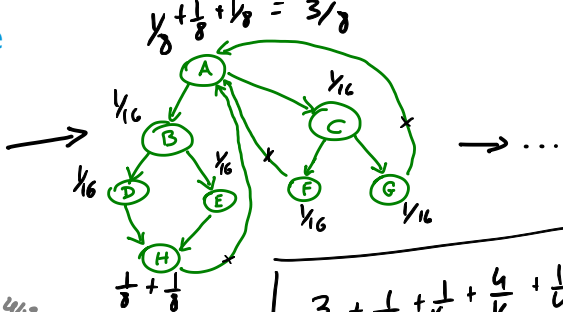
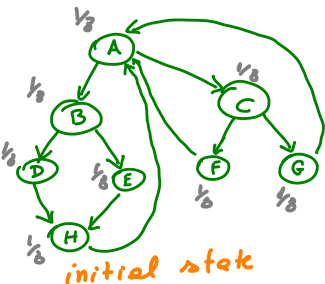
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# Page Rank - Example



$$\frac{3}{8} + \frac{1}{16} + \frac{1}{16} + \frac{4}{16} + \frac{1}{4}$$

$$\frac{6 + 1 + 1 + 4 + 4}{16} = 1$$

equilibrium



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Markov

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- **Equilibrium** of pagerank values equal to probabilities of **stationary distribution** of random walk

process

$$P \in \mathbb{R}^{n \times n}, \quad P_{i,j} = \frac{1}{\delta^{\text{out}}(j)}$$

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$$P \in \mathbb{R}^{n \times n}, \quad P_{i,j} = \frac{1}{\delta_{out}(j)}$$

- Pagerank values and transition probabilities satisfy same equations:

$$\underline{p_{t+1}(i)} = \sum_{i:(i,j) \in E} \underbrace{\frac{p_t(i)}{\delta_{out}(j)}}_{\substack{\text{rank at time } i \\ \# \text{ outgoing neighbors}}} \Rightarrow p_{t+1} = p_t \cdot P$$

$$\|p_{t+1}\|_1 = \|p_t\|_1 = 1$$

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- If graph finite, irreducible and aperiodic, fundamental theorem guarantees stationary distribution.

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  - Fix number  $0 < s < 1$
  - Divide  $s$  fraction of its pagerank value to its neighbors,
  - $1 - s$  fraction of its pagerank value to all nodes evenly

*making it sort of complete graph*

$\Delta G$  +  $(1-s)k_n$  *satisfies*  
*relevance* *FTMC*

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- Now resulting graph is *strongly connected* and *aperiodic*  $\Rightarrow$  unique stationary distribution

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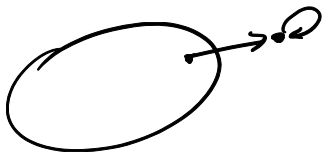
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- Now resulting graph is *strongly connected* and *aperiodic*  $\Rightarrow$  unique stationary distribution
- This modification does not change “relative importance” of vertices

# Conclusion

Markov Chains and Random Walks are ubiquitous in randomized algorithms.

- Page Rank algorithm (today's lecture)
- Approximation algorithms for counting problems [Karp, Luby & Madras]
  - Permanent of non-negative matrices [Jerrum, Vigoda & Sinclair]
- Sampling Problems
  - Gibbs sampling in statistical physics
  - many more places
- Probability amplification without too much randomness (efficient)
  - Random walks on expander graphs
- many more

# Potential Final Projects

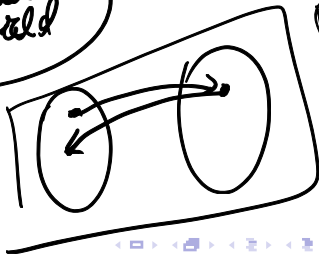


reducible  
(not strongly connected)

- Can we derandomize the perfect matching algorithms from class?
- A lot of progress has been made in the past couple years on this question in the works [?] and subsequently [?]
- Survey of the above, or understanding these papers is a great final project!

Conspiracy theory.

real world




periodic


# Acknowledgement

- Lecture based largely on:
  - Lap Chi's notes
  - [Motwani & Raghavan 2007, Chapter 6]
- See Lap Chi's notes at  
<https://cs.uwaterloo.ca/~lapchi/cs466/notes/L11.pdf>
- Also see Lap Chi's notes  
<https://cs.uwaterloo.ca/~lapchi/cs466/notes/L14.pdf> for a proof of fundamental theorem of Markov chains for undirected graphs.

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