

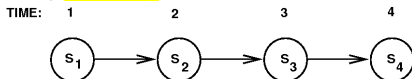
Lecture 8 - Reasoning under Uncertainty
(Part II)

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June 21, 2022

Readings: Poole & Mackworth (2nd ed.)Chapt. 8.5 - 8.9

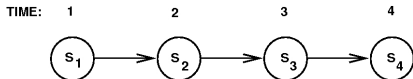
- A node repeats over **time**
- **explicit** encoding of time
- chain has length = amount of time you want to model
- **event-driven** times or **clock-driven** times
- e.g. **Markov chain**



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Markov assumption



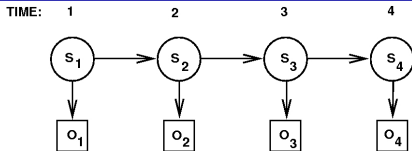
$$P(S_{t+1}|S_1, \dots, S_t) = P(S_{t+1}|S_t)$$

This distribution gives the **dynamics** of the Markov chain

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Hidden Markov Models (HMMs)



Add: observations O_t (always observed, so the node is square) and

observation function $P(O_t|S_t)$

Given a sequence of observations O_1, \dots, O_t , can estimate **filtering**:

$$P(S_t|O_1, \dots, O_t)$$

or **smoothing**, for $k < t$

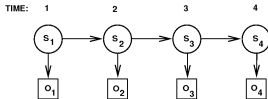
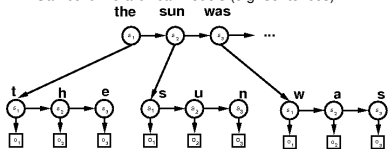
$$P(S_k|O_1, \dots, O_t)$$

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Speech Recognition

Belief Monitoring in HMMs

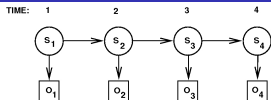
- Most well known application of HMMs
- **observations**: audio features
- **states**: phonemes
- **dynamics**: models e.g. co-articulation
- **HMMs**: words
- Can build hierarchical models (e.g. sentences)



filtering:

$$\begin{aligned}
 \alpha_j &= P(S_j | o_0, \dots, o_j) \\
 &\propto P(S_j, o_0, \dots, o_j) \\
 &= P(o_j | S_j) \sum_{S_{j-1}} P(S_j, S_{j-1}, o_0, \dots, o_{j-1}) \\
 &= P(o_j | S_j) \sum_{S_{j-1}} P(S_j | S_{j-1}) P(S_{j-1}, o_0, \dots, o_{j-1}) \\
 &\propto P(o_j | S_j) \sum_{S_{j-1}} P(S_j | S_{j-1}) \alpha_{j-1}
 \end{aligned}$$

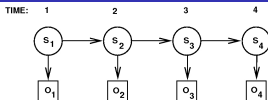
Belief Monitoring in HMMs



smoothing:

$$\begin{aligned}
 \beta_{i+1} &= P(o_{i+1}, \dots, o_T | S_i) \\
 &= \sum_{S_{i+1}} P(S_{i+1}, o_{i+1}, \dots, o_T | S_i) \\
 &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}, o_{i+2}, \dots, o_T, S_i) P(S_{i+1}, o_{i+2}, \dots, o_T | S_i) \\
 &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}) P(o_{i+2}, \dots, o_T | S_{i+1}, S_i) P(S_{i+1} | S_i) \\
 &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}) P(S_{i+1} | S_i) \beta_{i+2}
 \end{aligned}$$

Belief Monitoring in HMMs



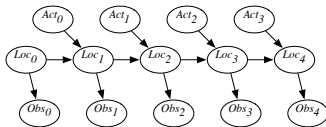
filtering and smoothing together:

$$\alpha_i \beta_{i+1} = P(o_{i+1}, \dots, o_T | S_i) P(S_i | o_0, \dots, o_i) \propto P(S_i | O)$$

- in general, any Bayesian network can repeat over time: **DBN**
- Many examples can be solved with **variable elimination**,
- may become too complex with enough variables
- **event-driven** times or **clock-driven** times



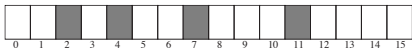
- Suppose a robot wants to determine its location based on its actions and its sensor readings: **Localization**
- This can be represented by the **augmented HMM** :



Example localization domain

Example Sensor Model

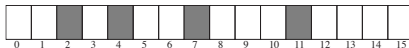
- **Circular** corridor, with 16 locations:



- Doors at positions: 2, 4, 7, 11.
- Noisy Sensors
- Stochastic Dynamics
- Robot starts at an unknown location and must determine where it is, known as the **kidnapped robot** problem.
- see handout `robot_loc.pdf`

- $P(\text{Observe Door} \mid \text{At Door}) = 0.8$
- $P(\text{Observe Door} \mid \text{Not At Door}) = 0.1$

- $P(Loc_{t+1} = l | Action_t = goRight \wedge Loc_t = l) = 0.1$
- $P(Loc_{t+1} = l + 1 | Action_t = goRight \wedge Loc_t = l) = 0.8$
- $P(Loc_{t+1} = l + 2 | Action_t = goRight \wedge Loc_t = l) = 0.074$
- $P(Loc_{t+1} = l' | Action_t = goRight \wedge Loc_t = l) = 0.002$ for any other location l' .
 - ▶ All location arithmetic is modulo 16.
 - ▶ The action *goLeft* works the same but to the left.



observe door, go right, observe no door, go right, observe door
 where is the robot?

$$P(Loc_2 = 4 | O_0 = d, A_0 = r, O_1 = \neg d, A_1 = r, O_2 = d) = 0.42$$

Combining sensor information

Probability Distribution and Monte Carlo

- **Example:** we can combine information from a light sensor and the door sensor **Sensor Fusion**
- **Key Point:** Bayesian probability ensures that evidence is integrated **proportionally to its precision.**
- Sensors are **precision weighted**



Loc_t robot location at time t
 D_t door sensor value at time t
 L_t light sensor value at time t



John von Neumann
1903 - 1957



ENIAC
1949



Stanislaw Ulam
1909-1984



Monte Carlo
1949

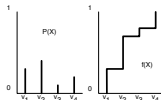
- **Idea:** probabilities \leftrightarrow samples
- Get probabilities from samples:

X	count		X	probability
x_1	n_1	\leftrightarrow	x_1	n_1/m
\vdots	\vdots		\vdots	\vdots
x_k	n_k		x_k	n_k/m
total	m			

- If we could **sample** from a variable's (posterior) probability, we could **estimate** its (posterior) probability.

For a variable X with a discrete domain or a (one-dimensional) real domain:

- **Totally order** the values of the domain of X .
- Generate the **cumulative probability distribution**: $f(x) = P(X \leq x)$.
- Select a value y **uniformly** in the range $[0, 1]$.
- Select the x such that $f(x) = y$.



Hoeffding Bound

p is true probability, s is sample average, n is number of samples

- $P(|s - p| > \epsilon) \leq 2e^{-2n\epsilon^2}$
- if we want an error greater than ϵ in less than a fraction δ of the cases, solve for n :

$$2e^{-2n\epsilon^2} < \delta$$

$$n > \frac{-\ln \frac{\delta}{2}}{2\epsilon^2}$$

- we have

ϵ error	cases with error $> \epsilon$	samples needed
0.1	1/20	184
0.01	1/20	18,445
0.1	1/100	265

Forward sampling in a belief network

- Sample the variables **one at a time**;
- sample **parents** of X **before** you sample X .
- Given values for the parents of X , sample from the **probability of X given its parents**.
- for samples $s_i, i = 1 \dots N$:

$$P(X = x_i) \propto \sum_{s_i} \delta(x_i) = N_{X=x_i}$$

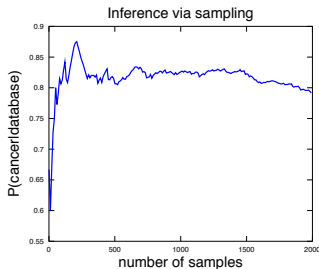
where

$$\delta(x_i) = \begin{cases} 1 & \text{if } X = x_i \text{ in } s_i \\ 0 & \text{otherwise} \end{cases}$$

Sample	Malfnction	Cancer	TestB	TestA	Report	Database
s_1	false	false	true	true	false	false
s_2	false	true	true	true	true	true
s_3	false	true	true	true	true	true
s_4	false	false	false	true	false	false
s_5	true	true	true	true	false	false
s_6	false	true	false	true	false	false
s_7	false	false	false	true	false	true
			...			
s_{1000}	false	false	false	true	false	false

To get $P(H = h_i | E = e_i)$ simply

- count the number of samples that have $H = h_i$ and $E = e_i$, $N(h_i, e_i)$
- divide by the number of samples that have $E = e_i$, $N(e_i)$
- $P(H = h_i | E = e_i) = \frac{P(H=h_i \wedge E=e_i)}{P(E=e_i)} = \frac{N(h_i, e_i)}{N(e_i)}$
- $P(C = True | Database = True)$ based on first 7 samples?



Rejection Sampling

Example Network

$\frac{P(A = true)}{0.4}$	$(A) \rightarrow (E)$	$\frac{A \quad P(E = true)}{true \quad 0.1}$
		$false \quad 0.3$

- To estimate a posterior probability given evidence $Y_1 = v_1 \wedge \dots \wedge Y_j = v_j$:
- If, for any i , a sample assigns Y_i to any value other than v_i **reject that sample**.
- The **non-rejected** samples are distributed according to the posterior probability.
- in the Hoeffding bound, n is the number of **non-rejected samples**

If we draw N samples $s_{i=1 \dots N}$ by

- sampling A : $a_{i=1 \dots N}$
- sampling from E given A : $e_{i=1 \dots N}$

then

- $\approx N_i = 0.4N$ of them will have $A = true$, and of these $\approx 10\%$ will have $E = true$
- $\approx N_i = 0.6N$ of them will have $A = false$, and of these $\approx 30\%$ will have $E = true$

$$\frac{P(A = \text{true})}{0.4}$$



A	P(E = true)
true	0.1
false	0.3

$$\frac{P(A = \text{true})}{0.4}$$



A	P(E = true)
true	0.1
false	0.3

so we have

A	E	N_{AE}
true	false	$N_{ff} = 0.4 \times 0.9 \times N$
true	true	$N_{tt} = 0.4 \times 0.1 \times N$
false	false	$N_{ff} = 0.6 \times 0.7 \times N$
false	true	$N_{ft} = 0.6 \times 0.3 \times N$

We want to compute

$$P(a|e) = P(A = \text{true}|E = \text{true}) \propto \sum_{s_i} \delta(a_i = \text{true}) \delta(e_i = \text{true})$$

$$\begin{aligned} P(a|e) &= \frac{P(a \wedge e)}{P(e)} \approx \frac{N_{tt}}{N_{tt} + N_{ft}} \\ &= \frac{0.1 \times 0.4 \times N}{0.1 \times 0.4 \times N + 0.3 \times 0.6 \times N} = 0.182 \end{aligned}$$

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- we can do better since we can **weight** the samples
- weights = prob. that the evidence is observed**
- N_t samples with $A = \text{true}$ have weight of $w_t = 0.1$ this is $P(E = \text{true}|A = \text{true})$
- N_f samples with $A = \text{false}$ have weight of $w_f = 0.3$ this is $P(E = \text{true}|A = \text{false})$
- can do better because we don't need to generate the 90% of samples (when $A = \text{true}$) that don't agree with the evidence - we simply assign all samples a weight of 0.1
- thus, we are **mixing exact inference** (the 0.1) with **sampling**.

$$\frac{P(A = \text{true})}{0.4}$$



A	P(E = true)
true	0.1
false	0.3

$$\frac{P(A = \text{true})}{0.4}$$



A	P(E = true)
true	0.1
false	0.3

- Compute sum of all weights of the samples with $A = \text{true}$

$$W_t = \sum_i w_i \delta(a_i = \text{true}) = N_t \times 0.1$$

- Compute sum of all weights of the samples with $A = \text{false}$

$$W_f = \sum_i w_i \delta(a_i = \text{false}) = N_f \times 0.3$$

- finally, compute

$$P(a|e) = \frac{W_t}{W_t + W_f} = \frac{0.1 \times 0.4 \times N}{0.1 \times 0.4 \times N + 0.3 \times 0.6 \times N}$$

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- In fact, the As don't need to even be sampled from $P(A)$
- Can be **sampled from some $q(A)$** , say $q(A = \text{true}) = 0.5$
- and each sample will have a new weight $P(a)/q(a)$
- $q(A)$ is a **proposal** distribution.
- helps when it is hard to sample from $P(A)$, but we can evaluate $P^*(A) \propto P(A)$ given a sample (see slide 24)

$\frac{P(A = \text{true})}{0.4}$	$(A) \rightarrow (E)$	<table border="1"> <tr> <td>A</td> <td>P(E = true)</td> </tr> <tr> <td>true</td> <td>0.1</td> </tr> <tr> <td>false</td> <td>0.3</td> </tr> </table>	A	P(E = true)	true	0.1	false	0.3
A	P(E = true)							
true	0.1							
false	0.3							

- rejection sampling uses $q = P$
- rejection sampling uses all variables including observed ones, and all weights on samples are set to 1.0
- $N_i^t = q(a)N$ samples with $A = \text{true}$ have weight of $0.1 \times \frac{P^*(a)}{q(a)} = 0.1 \times \frac{\alpha 0.4}{0.5}$
- $N_i^f = q(\bar{a})N$ samples with $A = \text{false}$ have weight of $0.3 \times \frac{P^*(\bar{a})}{q(\bar{a})} = 0.3 \times \frac{\alpha 0.6}{0.5}$

$\frac{P(A = \text{true})}{0.4}$	$(A) \rightarrow (E)$	<table border="1"> <tr> <td>A</td> <td>P(E = true)</td> </tr> <tr> <td>true</td> <td>0.1</td> </tr> <tr> <td>false</td> <td>0.3</td> </tr> </table>	A	P(E = true)	true	0.1	false	0.3
A	P(E = true)							
true	0.1							
false	0.3							

- total weight of all samples with $A = \text{true}$

$$W_i^t = \sum_i w_i \delta(a_i = \text{true}) = N_i^t \times 0.1 \times \frac{\alpha 0.4}{0.5}$$

$$= 0.5N \times 0.1 \times \frac{\alpha 0.4}{0.5} = 0.1 \times \alpha \times 0.4 \times N$$

- total weight of all samples with $A = \text{false}$

$$W_i^f = \sum_i w_i \delta(a_i = \text{false}) = N_i^f \times 0.3 \times \frac{\alpha 0.6}{0.5}$$

$$= 0.5N \times 0.3 \times \frac{\alpha 0.6}{0.5} = 0.3 \times \alpha \times 0.6 \times N$$

$\frac{P(A = \text{true})}{0.4}$	$(A) \rightarrow (E)$	<table border="1"> <tr> <td>A</td> <td>P(E = true)</td> </tr> <tr> <td>true</td> <td>0.1</td> </tr> <tr> <td>false</td> <td>0.3</td> </tr> </table>	A	P(E = true)	true	0.1	false	0.3
A	P(E = true)							
true	0.1							
false	0.3							

- finally, compute

$$P(a|e) = \frac{W_i^t}{W_i^t + W_i^f} = \frac{0.1 \times \alpha \times 0.4 \times N}{0.1 \times \alpha \times 0.4 \times N + 0.3 \times \alpha \times 0.6 \times N}$$

- Why** is it hard to sample from $P(A)$?
- because you need to know $P^*(A) \propto P(A)$ **for all values** of A (to normalize properly)
- in **high dimensional spaces**, there can be a lot
- consider the task of measuring the average (or maximum) depth of this lake - how do you draw samples? You cannot miss the canyons!



Figure 29.3. A slice through a lake that includes some canyons.

- $P(A) = \frac{P^*(A)}{\sum_A P^*(A)}$
- \sum_A is potentially **intractable**
- consider if A is continuous or discrete-valued over 500 dimensions.
- causes problems **for exact inference** as well

- Sometimes we may want to choose a proposal distribution that is **different** than the actual probability distribution
- We may want to **skew** the proposal - because we may have some **additional knowledge** about the data, for example
- or, we can generate proposals **from the data itself** using some procedural knowledge that is not directly encoded in the BN
- Can be important in multiple/many dimensions,

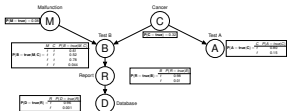
Recall variable elimination: To compute $P(Z, Y_1 = v_1, \dots, Y_j = v_j)$, we sum out the other variables, $Z_1, \dots, Z_k = \{X_1, \dots, X_n\} - \{Z\} - \{Y_1, \dots, Y_j\}$.

$$P(Z, Y_1 = v_1, \dots, Y_j = v_j) = \sum_{Z_k} \dots \sum_{Z_1} \prod_{i=1}^n P(X_i | \text{parents}(X_i))_{Y_1 = v_1, \dots, Y_j = v_j}$$

Now, we sample Z_{i+1}, \dots, Z_k and sum Z_1, \dots, Z_i ,

$$= \sum_{z_{i+1}, \dots, z_{k,j}} \left[\sum_{Z_1, \dots, Z_i} \prod_{j=1}^i P(Z_j | \text{parents}(Z_j))_{Y_1 = v_1, \dots, Y_j = v_j} \right] \frac{P(Z_{i+1}, \dots, Z_{k,j})}{q(Z_{i+1}, \dots, Z_{k,j})}$$

Importance Sampling example

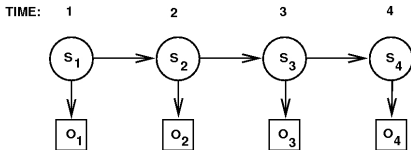


Compute $P(B|D = \text{true}, A = \text{false})$ by sampling C and M .

- use $q(C = \text{true}) = P(C = \text{true}) = 0.32$ and $q(M = \text{true}) = P(M = \text{true}) = 0.08$
- use $q(C = \text{true}) = 0.5$ and $q(M = \text{true}) = P(M = \text{true}) = 0.08$
- use $q(C = \text{true}) = q(M = \text{true}) = 0.5$

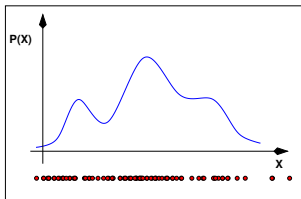
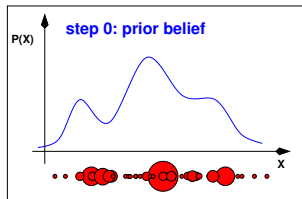
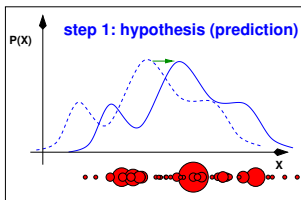
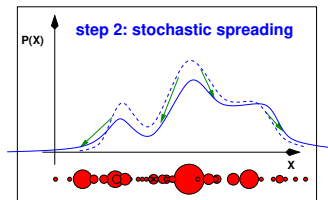
$$P(B|D = \text{true}, A = \text{false}) \propto \sum_{s_i = \{c_i, m_i\}} P(B, D = \text{true}, A = \text{false} | c_i, m_i)$$

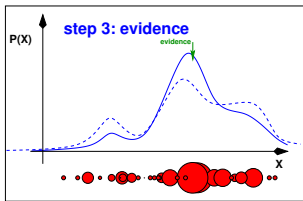
Stochastic Sampling for HMMs (and other DBNS)



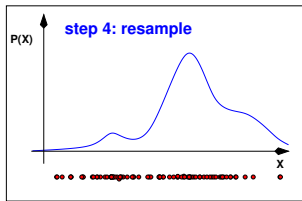
Sequential Monte Carlo or Particle Filter

- sequential stochastic sampling
- keep track of $P(S_t)$ at the current time t
- represent $P(S_t)$ with a set of samples
- update as new observations o_{t+1} arrive
 - predict $P(S_{t+1}) \propto P(S_{t+1}|S_t)$
 - compute weights as $P(o_{t+1}|S_{t+1})$
 - resample according to weights

sample i : $\{x_i\}$ sample i : $\{x_i, w_i\}$ sample i : $\{x_i, w_i\}$ sample i : $\{x_i, w_i\}$



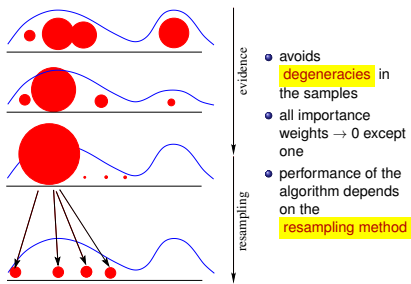
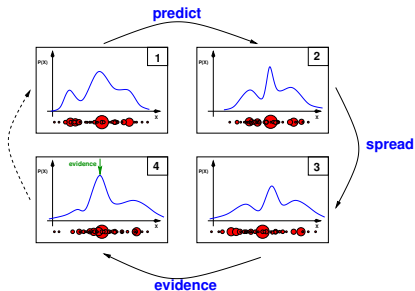
sample $i: \{x_i, w_i\}$



sample $i: \{x_i, w_i\}$

Bayesian Sequential Updates

Resampling



Next:

- Supervised Learning under Uncertainty (Poole & Mackworth (2nd ed.)10.1,10.4)