



Mode and Edgeworth Expansion for the Ewens Distribution and the Stirling Numbers

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Abstract

We provide asymptotic expansions for the Stirling numbers of the first kind and, more generally, the Ewens (or Karamata-Stirling) distribution. Based on these expansions, we obtain some new results on the asymptotic properties of the mode and the maximum of the Stirling numbers and the Ewens distribution. For arbitrary $\theta > 0$ and for all sufficiently large $n \in \mathbb{N}$, the unique maximum of the Ewens probability mass function

$$\mathbb{L}_n(k) = \frac{\theta^k}{\theta(\theta+1)\cdots(\theta+n-1)} \begin{bmatrix} n \\ k \end{bmatrix}, \quad k = 1, \dots, n,$$

is attained at $k = \lfloor a_n \rfloor$ or $\lceil a_n \rceil$, where $a_n = \theta \log n - \theta \Gamma'(\theta) / \Gamma(\theta) - 1/2$. We prove that the mode is the nearest integer to a_n for a set of n 's of asymptotic density 1, yet this formula is not true for infinitely many n 's.

1 Introduction and statement of results

1.1 Introduction

The (unsigned) *Stirling numbers* of the first kind $\begin{bmatrix} n \\ k \end{bmatrix}$ are defined, for $n \in \mathbb{N}$ and $1 \leq k \leq n$, by the formula

$$x^{(n)} := x(x+1)\cdots(x+n-1) = \sum_{k=1}^n \begin{bmatrix} n \\ k \end{bmatrix} x^k, \quad x \in \mathbb{R}. \quad (1)$$

For $n \in \mathbb{N}$, a random variable $K_n(\theta)$ is said to have the *Ewens distribution* with parameter $\theta > 0$ if its probability mass function is given by the formula

$$\mathbb{P}(K_n(\theta) = k) = \frac{\theta^k}{\theta^{(n)}} \begin{bmatrix} n \\ k \end{bmatrix}, \quad k = 1, \dots, n.$$

Bingham [2] called this distribution the *Karamata-Stirling law*. One can interpret $K_n(\theta)$ as the number of blocks in a random partition of $\{1, \dots, n\}$ distributed according to the Ewens sampling formula, or, equivalently, the number of different alleles in the infinite alleles model, the number of tables in a Chinese restaurant process, or the number of colors in the Hoppe urn. The Ewens sampling formula plays an important role in population genetics [6], [4, Section 1.3]. There is a distributional representation of $K_n(\theta)$ as a sum of independent random variables

$$K_n(\theta) \stackrel{d}{=} \xi_1 + \cdots + \xi_n, \quad \text{where } \xi_i \sim \text{Bern}(\theta/(\theta+i-1))$$

and $\text{Bern}(p)$ denotes the Bernoulli distribution with parameter p . In the special case $\theta = 1$, classical results going back at least to Feller [7] and Rényi [21] state that the random variable $K_n(1)$ has the same distribution as the number of cycles in a uniformly chosen random permutation of n objects, or the number of records in a sample of n i.i.d. variables from a

continuous distribution. It follows easily from Lindeberg's theorem that $K_n(\theta)$ satisfies a central limit theorem of the form

$$\frac{K_n(\theta) - \theta \log n}{\sqrt{\theta \log n}} \xrightarrow[n \rightarrow \infty]{d} \mathbf{N}(0, 1)$$

known as Goncharov's CLT in the case $\theta = 1$.

Asymptotic expansions, as $n \rightarrow \infty$, of the Stirling numbers $\begin{bmatrix} n \\ k \end{bmatrix}$ in various regions of k were provided in numerous works [12, 13, 18, 20, 23, 24]. Most notably, Hwang [13, Theorem 2] (and Theorem 14 on page 108 of his dissertation [12] for a more general result) gave an asymptotic expansion valid uniformly in the domain $2 \leq k \leq \eta \log n$, for any fixed $\eta > 0$. Louchard [18, Theorem 2.1] computed three non-trivial terms of the asymptotic expansion in the central regime $k = \log n + O(\sqrt{\log n})$ which is similar to the classical Edgeworth expansion in the central limit theorem.

In this short note we start by deriving a full Edgeworth expansion, as $n \rightarrow \infty$, for the sequence of probability mass functions $k \mapsto \mathbb{P}(K_n(\theta) = k)$ which is uniform both in $\theta \in [1/\eta, \eta]$ (where $\eta > 1$) and in $k \in \{1, \dots, n\}$; see Theorem 1. Our result is an application of the general Edgeworth expansion for deterministic or random profiles which the authors [16] recently obtained. Using this asymptotic expansion we derive some new results on the mode and the maximum of the Ewens distribution. In the case $\theta = 1$ the mode can be interpreted as the most probable number of cycles in a random permutation of n objects. It was investigated in the works of Hammersley [10, 11] and Erdős [5]. Our results on the mode and the maximum will be stated in Theorems 5 and 7 below.

1.2 Asymptotic expansion of the Ewens distribution

Before stating our main result we need to recall some notions. The (complete) *Bell polynomials* $B_j(z_1, \dots, z_j)$ are defined by the formal identity

$$\exp \left(\sum_{j=1}^{\infty} \frac{x^j}{j!} z_j \right) = \sum_{j=0}^{\infty} \frac{x^j}{j!} B_j(z_1, \dots, z_j).$$

Therefore $B_0 = 1$ and, for $j \in \mathbb{N}$,

$$B_j(z_1, \dots, z_j) = \sum' \frac{j!}{i_1! \cdots i_j!} \left(\frac{z_1}{1!} \right)^{i_1} \cdots \left(\frac{z_j}{j!} \right)^{i_j}, \quad (2)$$

where the sum \sum' is taken over all $i_1, \dots, i_j \in \mathbb{N}_0$ satisfying $1i_1 + 2i_2 + \cdots + ji_j = j$. For example, the first three Bell polynomials are given by

$$B_1(z_1) = z_1, \quad B_2(z_1, z_2) = z_1^2 + z_2, \quad B_3(z_1, z_2, z_3) = z_1^3 + 3z_1z_2 + z_3. \quad (3)$$

Further, we will use the ‘‘probabilist’’ *Hermite polynomials* $\text{He}_n(x)$ defined by

$$\text{He}_n(x) = e^{\frac{1}{2}x^2} \left(-\frac{d}{dx} \right)^n e^{-\frac{1}{2}x^2}, \quad n \in \mathbb{N}_0. \quad (4)$$

The first few Hermite polynomials needed for the first three terms of the expansion are

$$\begin{aligned} \text{He}_0(x) &= 1, & \text{He}_1(x) &= x, & \text{He}_2(x) &= x^2 - 1, & \text{He}_3(x) &= x^3 - 3x, \\ \text{He}_4(x) &= x^4 - 6x^2 + 3, & \text{He}_6(x) &= x^6 - 15x^4 + 45x^2 - 15. \end{aligned}$$

Theorem 1. Fix $r \in \mathbb{N}_0$ and a compact subset $L \subset (0, \infty)$. Uniformly over $\theta \in L$ we have

$$\lim_{n \rightarrow \infty} (\log n)^{\frac{r+1}{2}} \sup_{k=1, \dots, n} \left| \mathbb{P}(K_n(\theta) = k) - \frac{e^{-\frac{1}{2}x_n^2(k, \theta)}}{\sqrt{2\pi\theta \log n}} \sum_{j=0}^r \frac{H_j(x_n(k, \theta))}{(\theta \log n)^{j/2}} \right| = 0.$$

Here, $x_n(k, \theta) = \frac{k - \theta \log n}{\sqrt{\theta \log n}}$ and $H_j(x)$ is a polynomial of degree $3j$ given by

$$H_j(x) := H_j(x, \theta) = \frac{(-1)^j}{j!} e^{\frac{1}{2}x^2} B_j(\widetilde{D}_1, \dots, \widetilde{D}_j) e^{-\frac{1}{2}x^2}, \quad (5)$$

where B_j is the j -th Bell polynomial and $\widetilde{D}_1, \widetilde{D}_2, \dots$ are differential operators given by

$$\widetilde{D}_j := \widetilde{D}_j(\theta) = \frac{1}{(j+1)(j+2)} \left(\frac{d}{dx} \right)^{j+2} + \widetilde{\chi}_j(0) \left(\frac{d}{dx} \right)^j \quad (6)$$

with $\widetilde{\chi}_j(\beta) = -\left(\frac{d}{d\beta} \right)^j \log \Gamma(\theta e^\beta)$ and Γ denoting the Euler gamma function.

Remark 2. It follows from (3), (5) and (6) that the first three coefficients of the expansion are

$$\begin{aligned} H_0(x) &= 1, \\ H_1(x) &= -\frac{\Gamma'(\theta)}{\Gamma(\theta)} \theta x + \frac{1}{6} \text{He}_3(x), \\ H_2(x) &= \left(\theta^2 \frac{\Gamma'^2(\theta)}{\Gamma^2(\theta)} - \frac{\theta^2 \Gamma''(\theta) + \theta \Gamma'(\theta)}{2\Gamma(\theta)} \right) \text{He}_2(x) + \left(\frac{1}{24} - \frac{\Gamma'(\theta)}{6\Gamma(\theta)} \theta \right) \text{He}_4(x) + \frac{1}{72} \text{He}_6(x). \end{aligned}$$

An expression for $\widetilde{\chi}_j(0)$ involving polygamma functions and Stirling numbers of the second kind will be given below in (13). The tilde in \widetilde{D}_j and $\widetilde{\chi}_j$ is needed to keep the notation consistent with our more general work [16]. It is easy to check that $H_j(-x) = (-1)^j H_j(x)$ [16, Remark 2.4].

To compute $H_j(x)$ one can proceed as follows. First, express $\frac{1}{j!} B_j(\widetilde{D}_1, \dots, \widetilde{D}_j)$ as a polynomial in $D := \frac{d}{dx}$ (and note that only even/odd powers of D are present if j is even/odd). Then replace each occurrence of D^l by $\text{He}_l(x)$; see (4) for justification.

Remark 3. It is possible to choose the value of θ as a function of k . One natural choice is $\theta = 1$ which provides a full version of Louchard's expansion [18, Theorem 2.1] (although he used a slightly different normalization in his analogue of $x_n(k, 1)$ and his term $-355x^3/144$ should be

replaced by $-47x^3/144$). Another possible choice is $\theta = k/\log n$ (so that $x_n(k, \theta) = 0$), which gives a large-deviation-type expansion valid uniformly in the region $\eta^{-1} \log n < k < \eta \log n$, for fixed $\eta > 1$ and $q \in \mathbb{N}_0$:

$$\frac{(k/\log n)^k}{(k/\log n)^{(n)}} \left[\begin{matrix} n \\ k \end{matrix} \right] = \frac{1}{\sqrt{2\pi k}} \sum_{s=0}^q \frac{H_{2s}(0, k/\log n)}{k^s} + o\left(\frac{1}{(\log n)^{q+1}}\right).$$

Observe that the terms with half-integer powers of k are not present in the sum because $H_{2j+1}(0) = 0$. Using the formula

$$\frac{\Gamma(n + \theta)}{n!} = n^{\theta-1} \left(1 + O\left(\frac{1}{n}\right) \right)$$

yields the expansion

$$\frac{1}{n!} \left[\begin{matrix} n \\ k \end{matrix} \right] = \frac{1}{\Gamma(\theta)} n^{\theta - \theta \log \theta - 1} \left(\frac{1}{\sqrt{2\pi k}} \sum_{s=0}^q \frac{H_{2s}(0, \theta)}{k^s} + o\left(\frac{1}{(\log n)^{q+1}}\right) \right) \quad (7)$$

valid as $n \rightarrow \infty$ uniformly over k in the region $\theta = k/\log n \in (\eta^{-1}, \eta)$. In this region, this expansion must be equivalent to Hwang's result [13, Theorem 2]. It is not easy to rigorously verify this equivalence by a direct comparison, but we checked using Mathematica 9 that the first three non-trivial terms coincide. Note a misprint in the formula for the remainder term $Z_\mu(m, n)$ in Hwang [13, Theorem 2]: $(\log n)^m/(m!n)$ should be replaced by $(\log n)/(mn)$. Expansion (7) could be also deduced from the work of Féray et al. [8, Theorem 3.4].

Taking sums over k in Theorem 1 and using the Euler-Maclaurin formula to approximate Riemann sums by integrals, one obtains that

$$\mathbb{P}\left(\frac{K_n(\theta) - \theta \log n}{\sqrt{\theta \log n}} \leq x\right) = \Phi(x) + \frac{e^{-x^2/2}}{\sqrt{2\pi\theta \log n}} \left(\frac{1}{2} - \frac{x^2 - 1}{6} + \theta \frac{\Gamma'(\theta)}{\Gamma(\theta)} \right) + O\left(\frac{1}{\log n}\right),$$

uniformly in $x \in (\theta \log n)^{-1/2}(\mathbb{Z} - \theta \log n)$, where $\Phi(x)$ is the standard normal distribution function. The proof follows Grübel and Kabluchko [9, Proposition 2.5] and is therefore omitted. Yamato [25] recently stated a slightly incorrect version of this expansion missing the term $1/2$ which comes from the Euler-Maclaurin formula. Similarly, one can obtain further terms in the expansion of the distribution function of $(K_n(\theta) - \theta \log n)/\sqrt{\theta \log n}$.

Remark 4. Since the set $L \subseteq (0, \infty)$ in Theorem 1 has to be chosen compact, our results do not yield asymptotic expansions for $\mathbb{P}(K_n(\theta) = k)$ in the regime $k = o(\log n)$ of the same precision as Hwang's [13, Theorems 1 and 2]. Also, they do not extend straightforwardly to the case $k = n - O(n^\alpha)$ for $0 < \alpha < 1$ treated by Louchard [18, Section 3]. A generalization of our approach to cover these regions will be content of future work.

1.3 Mode and maximum of the Ewens distribution

Theorem 1 allows us to deduce various results on the *mode* and the *maximum* of the Ewens distribution. A mode is any value $k \in \{1, \dots, n\}$ maximizing $\mathbb{P}(K_n(\theta) = k)$, while the maximum $M_n(\theta)$ is defined by

$$M_n(\theta) = \max_{1 \leq k \leq n} \mathbb{P}(K_n(\theta) = k).$$

Let $u_n(\theta)$ denote the least mode. In this context, it is important to note that, for all $\theta > 0$, the function $k \mapsto \mathbb{P}(K_n(\theta) = k)$ is log-concave by a theorem attributed to Newton [11, 22], and

$$\begin{aligned} \mathbb{P}(K_n(\theta) = 1) &< \dots < \mathbb{P}(K_n(\theta) = u_n(\theta)) \\ &\geq \mathbb{P}(K_n(\theta) = u_n(\theta) + 1) > \dots > \mathbb{P}(K_n(\theta) = n). \end{aligned} \quad (8)$$

In particular, there are at most two modes. For $\theta = 1$, Erdős [5], proving a conjecture of Hammersley [11], showed that the mode is unique for all $n \geq 3$. By (8), uniqueness also holds for irrational θ ; however, for rational θ , the mode need not be unique since, for example,

$$\frac{2}{3} \begin{bmatrix} 3 \\ 1 \end{bmatrix} = \left(\frac{2}{3}\right)^2 \begin{bmatrix} 3 \\ 2 \end{bmatrix} > \left(\frac{2}{3}\right)^3 \begin{bmatrix} 3 \\ 3 \end{bmatrix}.$$

Theorem 5. Fix $\theta > 0$. There exists $N_1 \in \mathbb{N}$ such that for $n \geq N_1$, $u_n(\theta)$ is the unique mode of the Ewens distribution with parameter θ . The mode $u_n(\theta)$ equals one of the numbers $\lfloor u_n^*(\theta) \rfloor$ or $\lceil u_n^*(\theta) \rceil$, where

$$u_n^*(\theta) = \theta \log n - \frac{\theta \Gamma'(\theta)}{\Gamma(\theta)} - \frac{1}{2}$$

and $\lfloor \cdot \rfloor$, $\lceil \cdot \rceil$ denote the floor and the ceiling functions, respectively. Write $\delta_n(\theta) := \min_{k \in \mathbb{Z}} |u_n^*(\theta) - k|$. For the maximum $M_n(\theta)$, we have

$$\sqrt{2\pi\theta \log n} M_n(\theta) = 1 + \frac{\theta(\log \Gamma)'(\theta) + \theta^2(\log \Gamma)''(\theta) + 1/12 - \delta_n^2(\theta)}{2\theta \log n} + o\left(\frac{1}{\log n}\right).$$

In the case $\theta = 1$, Hammersley [11] and Erdős [5] derived related results for the mode. Cramer [3] discusses statistical applications and Mező [19] provides an overview and generalizations. Theorem 5 states that the mode is one of the numbers $\lfloor \log n + \gamma - \frac{1}{2} \rfloor$ or $\lceil \log n + \gamma - \frac{1}{2} \rceil$, for sufficiently large n . In fact, this holds for all $n \in \mathbb{N}$.

Proposition 6. $u_n(1) \in \{\lfloor \log n + \gamma - \frac{1}{2} \rfloor, \lceil \log n + \gamma - \frac{1}{2} \rceil\}$ for all $n \in \mathbb{N}$.

The proof uses the following formula of Hammersley [11]:

$$u_n(1) = \left\lfloor \log n + \gamma + \frac{\zeta(2) - \zeta(3)}{\log n + \gamma - \frac{3}{2}} + \frac{h(n)}{(\log n + \gamma - \frac{3}{2})^2} \right\rfloor, \quad (9)$$

for some $-1.098011 < h(n) < 1.430089$. Hwang [12, Section 5.7.9] gives a more precise expansion. Erdős [5] observed that, for $n > 189$, Hammersley's formula implies that the mode is one of the numbers $\lfloor \log(n-1) + \frac{1}{2} \rfloor$ or $\lfloor \log(n-1) + 1 \rfloor$. Note that his $\Sigma_{n,s}$ equals $\lfloor \frac{n+1}{n+1-s} \rfloor$ and his $n - f(n)$ is $u_{n+1}(1) - 1$ in our notation.

The next theorem provides more precise information about the behavior of the mode. Recall that a set $A \subset \mathbb{N}$ has *asymptotic density* $\alpha \in [0, 1]$ if

$$\lim_{n \rightarrow \infty} \frac{\#(A \cap \{1, \dots, n\})}{n} = \alpha.$$

For $x \in \mathbb{R}$, let $\{x\} = x - \lfloor x \rfloor$ denote the fractional part of x . Let $\text{nint}(x)$ be the integer closest to x (if $\{x\} = 1/2$, we agree to take $\text{nint}(x) = \lceil x \rceil$). That is,

$$\text{nint}(x) := \arg \min_{k \in \mathbb{Z}} |x - k| = \left\lfloor x + \frac{1}{2} \right\rfloor.$$

Theorem 7. *Fix $\theta > 0$. The mode $u_n(\theta)$ of the Ewens distribution with parameter θ has the following properties:*

- (i) *there exists a constant $C_0 > 0$ such that, for all $n \in \mathbb{N}$ satisfying*

$$\left| \{u_n^*(\theta)\} - \frac{1}{2} \right| > \frac{C_0}{\log n},$$

the mode $u_n(\theta)$ equals

$$\text{nint}(u_n^*(\theta)) = \left\lfloor \theta \log n - \frac{\theta \Gamma'(\theta)}{\Gamma(\theta)} \right\rfloor;$$

- (ii) *there are arbitrarily long intervals of consecutive n 's for which $u_n(\theta) = \lceil u_n^*(\theta) \rceil$; similarly, there are arbitrarily long intervals of consecutive n 's for which $u_n(\theta) = \lfloor u_n^*(\theta) \rfloor$;*
 (iii) *the set of $n \in \mathbb{N}$ such that $u_n(\theta) = \text{nint}(u_n^*(\theta))$ has asymptotic density one;*
 (iv) *there are infinitely many $n \in \mathbb{N}$ such that $u_n(\theta) \neq \text{nint}(u_n^*(\theta))$.*

The proof of part (iv) uses five terms in the Edgeworth expansion, where the first two terms influence the form of $u_n^*(\theta)$, while the remaining terms are needed for technical reasons. The idea is that the formula $u_n(\theta) = \text{nint}(u_n^*(\theta))$ becomes wrong if the fractional part of $u_n^*(\theta)$ is slightly below $\frac{1}{2}$, so that higher order terms in the Edgeworth expansion decide which of the values $\lfloor u_n^*(\theta) \rfloor$ and $\lceil u_n^*(\theta) \rceil$ is the mode. Using even more terms in the expansion, it is possible to replace $u_n^*(\theta)$ by some more complicated expressions involving higher-order corrections in inverse powers of $\theta \log n$ [12, Section 5.7.9]. However, it seems that there is no formula of the form

$$u_n(1) = \text{nint} \left(\log n + a_0 + \frac{a_1}{\log n} + \dots + \frac{a_r}{(\log n)^r} \right)$$

which is valid for all sufficiently large n .

Finally, we would like to mention that one can easily obtain counterparts of the above results for the B - and D -analogues of Stirling numbers of the first kind. These are defined as the coefficients of $(x+1)(x+3)\cdots(x+2n-1)$ and $((x+1)(x+3)\cdots(x+2n-3))(x+n-1)$, respectively. They appear, for example, in the study of intrinsic volumes of Weyl chambers [14].

2 Proofs

Proof of Theorem 1. The proof follows from the general Edgeworth expansion for random or deterministic profiles [16, Theorem 2.1]. We consider the sequence of “profiles”

$$\mathbb{L}_n(k) := \mathbb{P}(K_n(\theta) = k) = \frac{\theta^k}{\theta^{(n)}} \begin{bmatrix} n \\ k \end{bmatrix} \mathbb{1}_{\{k \in \{1, \dots, n\}\}},$$

and define

$$w_n := \theta \log n, \quad \varphi(\beta) := e^\beta - 1, \quad (\beta_-, \beta_+) = \mathbb{R}, \quad \mathcal{D} = \{z \in \mathbb{C} : |\operatorname{Im} z| < \pi\}.$$

In order to apply [16, Theorem 2.1], we need to check Conditions A1–A4 given in the beginning of Section 2 of the cited paper. Note that

$$\begin{aligned} W_n(\beta) &:= e^{-\varphi(\beta)w_n} \sum_{k \in \mathbb{Z}} e^{\beta k} \mathbb{L}_n(k) = n^{-\theta(e^\beta - 1)} \sum_{k=1}^n e^{\beta k} \frac{\theta^k}{\theta^{(n)}} \begin{bmatrix} n \\ k \end{bmatrix} \\ &= n^{-\theta(e^\beta - 1)} \frac{(\theta e^\beta)^{(n)}}{\theta^{(n)}} = n^{-\theta(e^\beta - 1)} \frac{\Gamma(\theta e^\beta + n) \Gamma(\theta)}{\Gamma(\theta e^\beta) \Gamma(\theta + n)} \xrightarrow{n \rightarrow \infty} \frac{\Gamma(\theta)}{\Gamma(\theta e^\beta)} =: W_\infty(\beta) \end{aligned}$$

locally uniformly in $\beta \in \mathcal{D}$ with a rate of convergence which is polynomial in n^{-1} . Hence Conditions A1–A3 are satisfied. In order to check A4 it is enough to show that for every $a > 0$, $r \in \mathbb{N}$ and every compact subset K_1 of \mathbb{R}

$$\sup_{\beta \in K_1} \sup_{a \leq u \leq \pi} \left(n^{-\theta(e^\beta - 1)} \left| \frac{\Gamma(\theta e^{\beta+iu} + n) \Gamma(\theta)}{\Gamma(\theta + n) \Gamma(\theta e^{\beta+iu})} \right| \right) = o(\log^{-r} n), \quad n \rightarrow \infty.$$

But this easily follows from

$$\begin{aligned} &\sup_{\beta \in K_1} \left(n^{-\theta(e^\beta - 1)} \sup_{a \leq u \leq \pi} \left| \frac{\Gamma(\theta e^{\beta+iu} + n) \Gamma(\theta)}{\Gamma(\theta + n) \Gamma(\theta e^{\beta+iu})} \right| \right) \\ &\leq C \sup_{\beta \in K_1} \left(n^{-\theta(e^\beta - 1)} \sup_{a \leq u \leq \pi} \left| \frac{\Gamma(\theta e^{\beta+iu} + n)}{\Gamma(\theta + n)} \right| \right) \leq C_1 \sup_{\beta \in K_1} n^{\theta e^\beta (\cos a - 1)}, \end{aligned}$$

with constants C, C_1 depending on K_1, θ and a . Therefore, Theorem 2.1 of [16] is applicable for the Ewens distribution with arbitrary fixed $\theta > 0$. In particular, for $\theta = 1$, we obtain

$$(\log n)^{\frac{r+1}{2}} \sup_{\beta \in K} \sup_{1 \leq k \leq n} \left| \frac{\Gamma(e^\beta) e^{\beta k}}{n^{e^\beta - 1} n!} \begin{bmatrix} n \\ k \end{bmatrix} - \frac{e^{-\frac{1}{2} x_n^2(k, e^\beta)}}{\sqrt{2\pi e^\beta \log n}} \sum_{j=0}^r \frac{G_j(x_n(k, e^\beta); \beta)}{(\log n)^{j/2}} \right| \xrightarrow[n \rightarrow \infty]{} 0, \quad (10)$$

where K is a compact subset of \mathbb{R} , and the polynomials G_0, G_1, \dots are defined as in Theorem 2.1 of [16]: for $j \in \mathbb{N}_0$, we have

$$G_j(x; \beta) = \frac{(-1)^j}{j!} e^{\frac{1}{2} x^2} B_j(D_1, \dots, D_j) e^{-\frac{1}{2} x^2} \quad (11)$$

with the differential operators

$$D_j := D_j(\beta) = e^{-\frac{1}{2} \beta j} \left(\frac{1}{(j+1)(j+2)} \left(\frac{d}{dx} \right)^{j+2} + \chi_j(\beta) \left(\frac{d}{dx} \right)^j \right), \quad (12)$$

where

$$\chi_j(\beta) = - \left(\frac{d}{d\beta} \right)^j \log \Gamma(e^\beta).$$

Now, if $L \subseteq (0, \infty)$ is compact, then $K := \log L$ is compact in \mathbb{R} . Applying (10) with $K = \log L$ and $\beta = \log \theta \in K$, we obtain

$$(\log n)^{\frac{r+1}{2}} \sup_{\theta \in L} \sup_{1 \leq k \leq n} \left| \frac{\Gamma(\theta) \theta^k}{n^{\theta-1} n!} \begin{bmatrix} n \\ k \end{bmatrix} - \frac{e^{-\frac{1}{2} x_n^2(k, \theta)}}{\sqrt{2\pi \theta \log n}} \sum_{j=0}^r \frac{G_j(x_n(k, \theta); \log \theta)}{(\log n)^{j/2}} \right| \xrightarrow[n \rightarrow \infty]{} 0.$$

By Stirling's formula, uniformly in $\theta \in L$, $n \in \mathbb{N}$ and $1 \leq k \leq n$, we have

$$\frac{\Gamma(\theta) \theta^k}{n^{\theta-1} n!} \begin{bmatrix} n \\ k \end{bmatrix} = \frac{\theta^k}{\theta^{(n)}} \begin{bmatrix} n \\ k \end{bmatrix} (1 + O(n^{-1})) = \frac{\theta^k}{\theta^{(n)}} \begin{bmatrix} n \\ k \end{bmatrix} + O(n^{-1}).$$

We conclude the proof by noting that $G_j(x; \log \theta) = \theta^{-j/2} H_j(x)$ which follows directly from $\tilde{\chi}_j(0) = \chi_j(\log \theta)$. Indeed, by comparing (6) and (12), we obtain

$$D_j(\log \theta) = \theta^{-j/2} \widetilde{D}_j(\theta),$$

which implies that

$$B_j(D_1(\log \theta), \dots, D_j(\log \theta)) = \theta^{-j/2} B_j(\widetilde{D}_1(\theta), \dots, \widetilde{D}_j(\theta))$$

since $B_j(z_1, \dots, z_j)$ is a sum of terms of the form $c \cdot z_1^{i_1} z_2^{i_2} \cdots z_j^{i_j}$ with $1i_1 + 2i_2 + \cdots + ji_j = j$; see (2). Comparing (5) and (11), we obtain the required identity $G_j(x; \log \theta) = \theta^{-j/2} H_j(x)$.

To see that $\tilde{\chi}_j(0) = \chi_j(\log \theta)$, one can easily show by induction over $j \geq 1$ that, both

$$\chi_j(\beta) = - \sum_{\ell=1}^j \left\{ \begin{matrix} j \\ \ell \end{matrix} \right\} \psi^{(\ell-1)}(e^\beta) e^{\ell\beta},$$

and

$$\tilde{\chi}_j(\beta) = - \sum_{\ell=1}^j \left\{ \begin{matrix} j \\ \ell \end{matrix} \right\} \psi^{(\ell-1)}(\theta e^\beta) (\theta e^\beta)^\ell. \quad (13)$$

Here $\psi^{(j)}(x) = (\log \Gamma(x))^{(j+1)}$ denotes the polygamma function and $\left\{ \begin{matrix} n \\ k \end{matrix} \right\}$ is the Stirling number of the second kind satisfying the recurrence

$$\left\{ \begin{matrix} n+1 \\ k \end{matrix} \right\} = \left\{ \begin{matrix} n \\ k-1 \end{matrix} \right\} + k \left\{ \begin{matrix} n \\ k \end{matrix} \right\}, \quad 1 \leq k \leq n, \quad n \in \mathbb{N},$$

with initial conditions $\left\{ \begin{matrix} 0 \\ 0 \end{matrix} \right\} = 1$, $\left\{ \begin{matrix} n \\ 0 \end{matrix} \right\} = \left\{ \begin{matrix} 0 \\ n \end{matrix} \right\} = 0$. □

Proof of Theorem 5. It follows from Theorems 2.10 in [16] that for sufficiently large n , the maximizers of the function $k \mapsto \mathbb{P}(K_n(\theta) = k)$ must be of the form $\lfloor u_n^* \rfloor$ or $\lceil u_n^* \rceil$.

Next we prove that the maximizer is unique (for sufficiently large n) by following a method of Erdős [5] who considered the case $\theta = 1$. Thanks to (8), the uniqueness is evident if θ is irrational. Hence, we assume that $\theta = Q_1/Q_2$ is rational with Q_1, Q_2 being integer. We have, by (1),

$$\left[\begin{matrix} n \\ k \end{matrix} \right] = \sum_{1 \leq a_1 < \dots < a_{n-k} \leq n-1} a_1 \cdots a_{n-k}.$$

Put $k_n = \lceil u_n^*(\theta) \rceil = \theta \log n + O(1)$ as $n \rightarrow \infty$. By (8), it is sufficient to show that

$$\theta^{k_n} \left[\begin{matrix} n \\ k_n \end{matrix} \right] \neq \theta^{k_n-1} \left[\begin{matrix} n \\ k_n-1 \end{matrix} \right]. \quad (14)$$

By Erdős' argument relying on the prime number theorem with an appropriate error term [5, p. 233], for all sufficiently large n , there is a prime number p satisfying $(n-1)/k_n < p < (n-1)/(k_n-1)$. Then,

$$\left[\begin{matrix} n \\ k_n \end{matrix} \right] \not\equiv 0 \pmod{p}, \quad \left[\begin{matrix} n \\ k_n-1 \end{matrix} \right] \equiv 0 \pmod{p}$$

because in the representation of the former Stirling number all products except one are divisible by p , whereas in the latter all products are divisible by p . If n is large, p is not among the prime factors of Q_1 and Q_2 . Hence (14) follows and the mode of $K_n(\theta)$ is unique. Finally, the formula for M_n follows from Theorem 2.13 of [16]. □

Proof of Proposition 6. Recall Hammersley's formula (9):

$$u_n(1) = \left[\log n + \gamma + \frac{\zeta(2) - \zeta(3)}{\log n + \gamma - \frac{3}{2}} + \frac{h(n)}{(\log n + \gamma - \frac{3}{2})^2} \right]$$

with some $-1.1 < h(n) < 1.44$. It is easy to check that

$$\frac{\zeta(2) - \zeta(3)}{x} - \frac{1.1}{x^2} > -\frac{1}{2} \text{ and } \frac{\zeta(2) - \zeta(3)}{x} + \frac{1.44}{x^2} < \frac{1}{2}$$

for $x > 2.5$. Hence, the proposition is true for $\log n + \gamma - \frac{3}{2} > 2.5$, that is for $n \geq 31$. For $n = 1, 2, \dots, 30$ the statement is easy to verify using Mathematica 9. \square

Proof of Theorem 7 (i) and (ii). Part (i) follows essentially from Theorem 2.10 in [16] and its proof. Namely, by [16, Equation (90)], for $k = k(n) = u_n^*(\theta) + g \in \mathbb{Z}$ with $g = O(1)$, we have

$$\sqrt{2\pi\theta \log n} (\mathbb{P}(K_n(\theta) = k + 1) - \mathbb{P}(K_n(\theta) = k)) = -\frac{2g + 1}{2\theta \log n} + o\left(\frac{1}{\log n}\right).$$

The same relation, but with a better remainder term $O(\frac{1}{\log^2 n})$, follows from (16) which we shall prove below. Taking $g = -\{u_n^*(\theta)\}$, so that $k = \lfloor u_n^*(\theta) \rfloor$ and $k + 1 = \lceil u_n^*(\theta) \rceil$, yields

$$\mathbb{P}(K_n(\theta) = \lceil u_n^*(\theta) \rceil) - \mathbb{P}(K_n(\theta) = \lfloor u_n^*(\theta) \rfloor) = \frac{1}{\sqrt{2\pi\theta \log n}} \left(\frac{\{u_n^*(\theta)\} - \frac{1}{2}}{\theta \log n} + O\left(\frac{1}{\log^2 n}\right) \right).$$

It follows that there is a sufficiently large constant $C_0 > 0$ such that, if $\{u_n^*(\theta)\} > \frac{1}{2} + \frac{C_0}{\log n}$, then the right-hand side is positive, and the mode equals $\lceil u_n^*(\theta) \rceil$. Similarly, if $\{u_n^*(\theta)\} < \frac{1}{2} - \frac{C_0}{\log n}$, then the right-hand side is negative, and the mode equals $\lfloor u_n^*(\theta) \rfloor$.

The proof of part (ii) follows immediately from part (i) and the fact that, for every fixed $L > 0$, we have $\log(n + L) - \log n \rightarrow 0$ as $n \rightarrow \infty$. \square

Proof of Theorem 7 (iii). In view of part (i) it suffices to show that

$$\limsup_{\varepsilon \rightarrow 0} \limsup_{n \rightarrow \infty} \frac{\#\{1 \leq k \leq n : \text{dist}(u_k^*(\theta), \mathbb{Z} + 1/2) < \varepsilon\}}{n} = 0,$$

which, in turn, follows from the fact that

$$\limsup_{\varepsilon \rightarrow 0} \limsup_{n \rightarrow \infty} \frac{\#\{1 \leq k \leq n : \text{dist}(\log k, \alpha\mathbb{Z} + \beta) < \varepsilon\}}{n} = 0, \quad (15)$$

for all $\alpha > 0$ and $\beta \in \mathbb{R}$. Equation (15) would be true if the sequence of fractional parts of $\alpha^{-1} \log k$, $k \in \mathbb{N}$, were uniformly distributed on $[0, 1]$. However, the latter claim is unfortunately not true [17, Examples 2.4 and 2.5, pp. 8–9]. Let us prove (15). We have, assuming

that $\varepsilon < \alpha/2$,

$$\begin{aligned}
\#\{1 \leq k \leq n : \text{dist}(\log k, \alpha\mathbb{Z} + \beta) < \varepsilon\} &= \sum_{k=1}^n \#\{j \in \mathbb{Z} : \text{dist}(\log k, \alpha j + \beta) < \varepsilon\} \\
&= \sum_{j \in \mathbb{Z}} \#\{1 \leq k \leq n : e^{\alpha j + \beta - \varepsilon} < k < e^{\alpha j + \beta + \varepsilon}\} \\
&\leq \sum_{j \in \mathbb{Z}} \#\{k \in \mathbb{N} : e^{\alpha j + \beta - \varepsilon} \vee 1 \leq k \leq e^{\alpha j + \beta + \varepsilon} \wedge n\}.
\end{aligned}$$

The summand on the right-hand side is the number of integers in the interval $[e^{\alpha j + \beta - \varepsilon} \vee 1, e^{\alpha j + \beta + \varepsilon} \wedge n]$ (which is empty if either $e^{\alpha j + \beta - \varepsilon} > n$ or $e^{\alpha j + \beta + \varepsilon} < 1$). Hence, it is bounded from above by $(e^{\alpha j + \beta + \varepsilon} \wedge n - e^{\alpha j + \beta - \varepsilon} \vee 1 + 1)_+$. Therefore,

$$\#\{1 \leq k \leq n : \text{dist}(\log k, \alpha\mathbb{Z} + \beta) < \varepsilon\} \leq \sum_{j \in \mathbb{Z}} (e^{\alpha j + \beta + \varepsilon} \wedge n - e^{\alpha j + \beta - \varepsilon} \vee 1 + 1)_+.$$

Further,

$$\begin{aligned}
&\sum_{j \in \mathbb{Z}} (e^{\alpha j + \beta + \varepsilon} \wedge n - e^{\alpha j + \beta - \varepsilon} \vee 1 + 1)_+ \\
&= \sum_{j \in \mathbb{Z}} e^{\alpha j + \beta + \varepsilon} \mathbb{1}_{\{\alpha j + \beta + \varepsilon < 0\}} + \sum_{j \in \mathbb{Z}} e^{\alpha j + \beta + \varepsilon} \mathbb{1}_{\{\alpha j + \beta - \varepsilon < 0, 0 \leq \alpha j + \beta + \varepsilon < \log n\}} \\
&\quad + \sum_{j \in \mathbb{Z}} n \mathbb{1}_{\{\alpha j + \beta - \varepsilon < 0, \log n \leq \alpha j + \beta + \varepsilon\}} \\
&\quad + \sum_{j \in \mathbb{Z}} (e^{\alpha j + \beta + \varepsilon} - e^{\alpha j + \beta - \varepsilon} + 1) \mathbb{1}_{\{\alpha j + \beta - \varepsilon \geq 0, \alpha j + \beta + \varepsilon < \log n\}} \\
&\quad + \sum_{j \in \mathbb{Z}} (n - e^{\alpha j + \beta - \varepsilon} + 1)_+ \mathbb{1}_{\{\alpha j + \beta - \varepsilon \geq 0, \log n \leq \alpha j + \beta + \varepsilon\}}.
\end{aligned}$$

Note that the first series converges, the second contains at most one summand since we assume $\varepsilon < \alpha/2$, and the third vanishes for n large enough. It can be checked that

$$\sum_{j \in \mathbb{Z}} (e^{\alpha j + \beta + \varepsilon} - e^{\alpha j + \beta - \varepsilon} + 1) \mathbb{1}_{\{\alpha j + \beta - \varepsilon \geq 0, \alpha j + \beta + \varepsilon < \log n\}} \leq C(\alpha, \beta)(e^{\beta + \varepsilon} - e^{\beta - \varepsilon})n$$

with an absolute constant $C(\alpha, \beta)$. Further, for n sufficiently large, we have

$$\sum_{j \in \mathbb{Z}} (n - e^{\alpha j + \beta - \varepsilon} + 1)_+ \mathbb{1}_{\{\alpha j + \beta - \varepsilon \geq 0, \log n \leq \alpha j + \beta + \varepsilon\}} \leq n(1 - e^{-2\varepsilon}) + 1.$$

Putting pieces together gives (15). □

Proof of Theorem 7 (iv). Recall the notation $w_n = \theta \log n$ and $x_n(k) = x_n(k, \theta) = (k - w_n)/\sqrt{w_n}$. Using Theorem 1 with $r = 4$, we obtain

$$\begin{aligned} \sqrt{2\pi w_n} \mathbb{P}(K_n(\theta) = k) &= e^{-\frac{1}{2}x_n^2(k)} \\ &\times \left(1 + \frac{H_1(x_n(k))}{w_n^{1/2}} + \frac{H_2(x_n(k))}{w_n} + \frac{H_3(x_n(k))}{w_n^{3/2}} + \frac{H_4(x_n(k))}{w_n^2} + o\left(\frac{1}{\log^2 n}\right) \right), \end{aligned}$$

as $n \rightarrow \infty$ uniformly in $1 \leq k \leq n$. Now let $k = \theta \log n + a$, where $a = O(1)$ as $n \rightarrow \infty$, so that $x_n(k) = a/w_n^{1/2}$. We have

$$\begin{aligned} H_1(x_n(k)) &= A_{11}(\theta) \frac{a}{w_n^{1/2}} + A_{12}(\theta) \frac{a^3}{w_n^{3/2}}, \\ H_2(x_n(k)) &= A_{21}(\theta) + A_{22}(\theta) \frac{a^2}{w_n} + o\left(\frac{1}{w_n}\right), \\ H_3(x_n(k)) &= A_{31}(\theta) \frac{a}{w_n^{1/2}} + o\left(\frac{1}{w_n^{1/2}}\right), \\ H_4(x_n(k)) &= A_{41}(\theta) + o(1), \end{aligned}$$

where $A_{11}(\theta), \dots, A_{41}(\theta)$ are some polynomials in $\widetilde{\chi}_1(0), \widetilde{\chi}_2(0), \widetilde{\chi}_3(0)$ and $\widetilde{\chi}_4(0)$; see Remark 2. Plugging these expressions into the asymptotic expansion above and using the expansion $e^y = 1 + y + y^2/2 + o(y^2)$, as $y \rightarrow 0$, yields

$$\sqrt{2\pi w_n} \mathbb{P}(K_n(\theta) = k) = 1 - \left(\frac{a^2}{2} - A_{11}(\theta)a - A_{21}(\theta) \right) \frac{1}{w_n} + \frac{P_\theta(a)}{w_n^2} + o\left(\frac{1}{\log^2 n}\right),$$

where

$$P_\theta(a) := \frac{1}{8}a^4 + \left(A_{12}(\theta) - \frac{1}{2}A_{11}(\theta) \right) a^3 + \left(A_{22}(\theta) - \frac{1}{2}A_{21}(\theta) \right) a^2 + A_{31}(\theta)a + A_{41}(\theta).$$

Now let us write $k = \theta \log n + a^* + g$, where $a^* := A_{11}(\theta) = -\frac{\theta\Gamma'(\theta)}{\Gamma(\theta)} - \frac{1}{2}$, yielding

$$\sqrt{2\pi w_n} \mathbb{P}(K_n(\theta) = k) = 1 - \left(\frac{g^2 - (a^*)^2}{2} - A_{21}(\theta) \right) \frac{1}{w_n} + \frac{P_\theta(a^* + g)}{w_n^2} + o\left(\frac{1}{\log^2 n}\right). \quad (16)$$

We are interested in g being either $\lfloor u_n^*(\theta) \rfloor - u_n^*(\theta) =: g'_n$ or $\lceil u_n^*(\theta) \rceil - u_n^*(\theta) =: g''_n$. Let M be the set of natural numbers n with $\{u_n^*(\theta)\} < 1/2 < \{u_{n+1}^*(\theta)\}$. Note that M has infinitely many elements because $\log n \rightarrow \infty$ and $\log(n+1) - \log n \rightarrow 0$. In the remainder of the proof, we always consider $n \in M$. Since $u_{n+1}^*(\theta) - u_n^*(\theta) = O(n^{-1})$, we have

$$g'_n = -1/2 + O(n^{-1}), \quad g''_n = 1/2 + O(n^{-1}).$$

Putting $k = \lfloor u_n^*(\theta) \rfloor$ into (16) yields

$$\begin{aligned} & \sqrt{2\pi w_n} \mathbb{P}(K_n(\theta) = \lfloor u_n^*(\theta) \rfloor) \\ &= 1 - \left(\frac{(g'_n)^2 - (a^*)^2}{2} - A_{21}(\theta) \right) \frac{1}{w_n} + \frac{P_\theta(a^* + g'_n)}{w_n^2} + o\left(\frac{1}{\log^2 n}\right) \\ &= 1 - \left(\frac{1 - 4(a^*)^2}{8} - A_{21}(\theta) \right) \frac{1}{w_n} + \frac{P_\theta(a^* - 1/2)}{w_n^2} + o\left(\frac{1}{\log^2 n}\right). \end{aligned}$$

Analogously, putting $k = \lceil u_n^*(\theta) \rceil$ gives

$$\sqrt{2\pi w_n} \mathbb{P}(K_n(\theta) = \lceil u_n^*(\theta) \rceil) = 1 - \left(\frac{1 - 4(a^*)^2}{8} - A_{21}(\theta) \right) \frac{1}{w_n} + \frac{P_\theta(a^* + 1/2)}{w_n^2} + o\left(\frac{1}{\log^2 n}\right).$$

For sufficiently large n the mode $u_n(\theta)$ equals either $\lfloor u_n^*(\theta) \rfloor$ or $\lceil u_n^*(\theta) \rceil$ depending on the sign of

$$s^*(\theta) := P_\theta(a^* + 1/2) - P_\theta(a^* - 1/2).$$

In the following we shall show that $s^*(\theta) > 0$, hence $u_n(\theta) = \lceil u_n^*(\theta) \rceil$, while $\text{nint}(u_n(\theta)) = \lfloor u_n^*(\theta) \rfloor$, so that $u_n(\theta) \neq \text{nint}(u_n^*(\theta))$. Recalling the polygamma function $\psi^{(m)}(\theta) = (\log \Gamma(\theta))^{(m+1)}$, the authors [15] checked with the help of Mathematica 9 that

$$s^*(\theta) = \frac{\theta^2}{2} (2\psi^{(1)}(\theta) + \theta\psi^{(2)}(\theta)).$$

Using the well-known formula for the polygamma function [1, 6.4.10]

$$\psi^{(m)}(\theta) = (\log \Gamma(\theta))^{(m+1)} = (-1)^{m+1} m! \sum_{k=0}^{\infty} \frac{1}{(\theta + k)^{m+1}}, \quad -\theta \notin \mathbb{N}_0, \quad m \geq 1,$$

we finally obtain

$$s^*(\theta) = \theta^2 \sum_{k=1}^{\infty} \frac{k}{(\theta + k)^3}, \quad \theta > 0,$$

yielding positivity of $s^*(\theta)$ for all $\theta > 0$. The proof of part (iv), as well as of the whole theorem, is complete. \square

Remark 8. For $\theta = 1$ we have $s^*(1) = \zeta(2) - \zeta(3)$, a term appearing in Hammersley's formula (9). In fact, in the special case $\theta = 1$ part (iv) could be deduced directly from (9).

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